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Scaling and merging time-resolved pink-beam diffraction with variational inference

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ABSTRACT

Time-resolved x-ray crystallography (TR-X) at synchrotrons and free electron lasers is a promising technique for recording dynamics of molecules at atomic resolution. While experimental methods for TR-X have proliferated and matured, data analysis is often difficult. Extracting small, time-dependent changes in signal is frequently a bottleneck for practitioners. Recent work demonstrated this challenge can be addressed when merging redundant observations by a statistical technique known as variational inference (VI). However, the variational approach to time-resolved data analysis requires identification of successful hyperparameters in order to optimally extract signal. In this case study, we present a successful application of VI to time-resolved changes in an enzyme, DJ-1, upon mixing with a substrate molecule, methylglyoxal. We present a strategy to extract high signal-to-noise changes in electron density from these data. Furthermore, we conduct an ablation study, in which we systematically remove one hyperparameter at a time to demonstrate the impact of each hyperparameter choice on the success of our model. We expect this case study will serve as a practical example for how others may deploy VI in order to analyze their time-resolved diffraction data.

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I. INTRODUCTION

Time-resolved x-ray crystallography (TR-X) enables the observation of macromolecular dynamics at atomic resolution, capturing essential information to understand reaction mechanisms. The applications of this technique are extremely broad, and within this paradigm, many experimental approaches have been developed at both synchrotrons and x-ray free electron lasers (XFELs). These include pump-probe,^{1–3} mix-and-inject serial crystallography (MISC),^{4–7} temperature jump,⁸ electric field stimulation,⁹ and others, which allow researchers to observe conformational changes

of diverse biological targets.^{10,11} In this manuscript, we focus on the data analysis of a successful MISC pink-beam diffraction experiment carried out at the Advanced Photon Source Sector 14 (BioCARS^{12,13}). This experiment investigated the reaction of a human enzyme implicated in early-onset Parkinson's disease, DJ-1, with one of its putative substrates, methylglyoxal. While this is a fascinating biochemical system of considerable medical relevance, here we will focus on the challenges inherent in extracting time-resolved signal from x-ray diffraction experiments rather than the biological

conclusions of our study. Based on our experience, we offer practical advice to those conducting time-resolved and comparative crystallography experiments.

The analysis of time-resolved diffraction data is still frequently challenging. The experiments themselves introduce many sources of variability. The goal of any time-resolved experiment is to initiate dynamics as uniformly as possible for all of the protein molecules inside a crystal. Then, assuming sufficient initiation and reaction progress, small structural changes occur and are captured at predetermined delay times with x-ray pulses. During data collection, the average structure in time across the crystal is recorded. Many time-resolved measurements are coupled to serial crystallography experiments, where datasets comprise up to hundreds of thousands of crystals, introducing additional challenges, such as heterogeneity across microcrystals and the need for accurate scaling and merging of reflections across the dataset. Time-resolved differences are most commonly analyzed with difference electron density (TR-DED) maps. The Fourier coefficients of such a map consist of the difference in structure factor amplitudes between two time points. The phases are usually approximated by those of a high-resolution reference structure of the crystal in the ground state. While difference electron density (DED) maps are exquisitely sensitive to changes in the composition of crystals,¹⁴ they are equally sensitive to measurement errors.¹⁵ The success of a time-resolved crystallography experiment may therefore hinge upon the accuracy with which differences in structure factor amplitudes can be measured, which is a noteworthy challenge. The changes in structure factor amplitudes between time points are, in most cases, small relative to the structure factors themselves.¹⁴ This means that data must be measured with unusually high precision in order to accurately infer differences in structure and observe intermediate states. Contamination of the structure factor differences by outliers can mask signal in otherwise well-measured data.¹⁵

Serial femtosecond crystallography (SFX) experiments were pioneered at x-ray free electron lasers,¹⁶ and serial synchrotron crystallography (SSX) experiments are becoming increasingly popular, with both monochromatic^{17–20} and polychromatic x rays.^{21,22} Data collection from quasi-monochromatic sources, however, records only partial Bragg reflections. The most common strategy to handle this partiality, and thus achieve more precise estimates of structure factors, is serial crystallography at extreme statistical redundancy.²³ Our experimental design is different and exploits a polychromatic synchrotron beamline (BioCARS 14-ID, Advanced Photon Source), which has already been used for previously successful SSX experiments.^{21,24,25} The increased reciprocal space information available in polychromatic diffraction patterns, as well as recording of full rather than partial intensities (without crystal rotation), as compared to monochromatic images enabled us to successfully measure time-resolved changes in structure from fewer microcrystals than are typically required in an SFX experiment. Other synchrotron beamlines are starting to broaden their x-ray bandwidth to $\sim 1\%$ due to these benefits.^{26–28} Nonetheless, many of our conclusions will be applicable to time-resolved crystallography experiments, regardless of the x-ray bandwidth or source. The central insight we wish to convey in this manuscript is that recent advances in data analysis can address the measurement errors inherent in crystallography experiments and generate high-quality DED maps. More specifically, we will demonstrate that the statistical method of variational inference^{29,30} is well suited to this challenge.

X-ray diffraction data typically contain many redundant observations of reflections. Merging is the process of averaging the intensities of redundant measurements to estimate a consensus set of intensities, which can be used in structure determination or, indeed, Fourier difference map analyses. This task is complicated by the nature of the diffraction experiment, as discussed above. The observed intensity is always modulated by a number of effects, which are unavoidable and lead to systematic errors in the measurements. These errors cannot be accounted for strictly by analytical correction factors.³¹ Prior to merging, observed intensities are corrected using numerical optimization, a procedure known as scaling.³² Conventionally, scaling algorithms work by fitting a model of systematic errors in order to minimize the discrepancy between repeated observations of the same reflection. This has been a very successful strategy for conventional, rotation-method data. Every major crystallography software package implements a scaling algorithm. An alternative to sequential scaling and merging was recently proposed, which jointly estimates merged structure factors alongside the systematic error model using variational inference.³³ This algorithm is implemented in the Python package, *careless*, which is freely available and open source (<https://github.com/rs-station/careless>).

Variational inference is a Bayesian method which uses gradient-based optimization to estimate a quantity of interest given some experimental observations and a set of prior beliefs.^{29,30} This is similar in concept to maximum a posterior (MAP) or regularized maximum-likelihood estimation. However, it differs in that a distribution is recovered for the estimand rather than a point estimate. Importantly, the width of this distribution can be used to assign a confidence to the estimate. In the context of *careless*, the estimand is the structure factor amplitudes, and the data are unmerged, unscaled intensities. The prior distribution is Wilson's priors,³⁴ which are derived under the assumption of random atomic coordinates within a crystalline lattice. Because variational inference relies on gradient-based optimization, it can be easily implemented in a modern deep-learning framework like PyTorch or TensorFlow,^{35,36} which provides automatic differentiation. In light of this, variational inference models frequently incorporate components based on deep learning. *Careless* is no exception. Indeed, a deep neural network is used to estimate systematic errors from experimental metadata. This is a black-box function, which ingests metadata for each reflection observation and outputs an estimate of the systematic error. In *careless*, parameters of this network are learned jointly with the structure factor amplitudes. We explore the benefits of this approach, specifically in the context of optimizing time-resolved difference peaks, in this work.

The main advantage of variational inference is flexibility. Consequently, software like *careless* should not be thought of as a push-button scaling solution. Rather, it is a modeling toolkit that can be used to extract the most signal possible from a time series. Doing so requires the user to identify the proper set of model hyperparameters to maximize the desired signal. Extensive hyperparameter searches can be computationally prohibitive. Therefore, we offer some lessons from our own experience to help decrease the burden of finding the optimal hyperparameters.

II. METHODS

A. Experimental design

All diffraction data were collected from DJ-1 microcrystals ($\sim 25 \mu\text{m}$ in size, grown as previously described^{5,37}) at APS BioCARS 14-ID-B^{12,13} with a custom sample cell coupled to a microfluidic

```

1  #!/bin/bash
2
3  R=0.99 #Multi wilson R-value
4  BASE_ARGS=(
5  --mc-samples=20
6  --test-fraction=0.05
7  --merge-half-datasets
8  --half-dataset-repeats=10
9  --dmin=1.77
10 --iterations=30_000
11 --separate-files
12 --double-wilson-parents=None,0,0,0,0,0
13 --double-wilson-r=0,$R,$R,$R,$R,$R,$R
14 --image-layers=2
15 --studentt-likelihood-dof=32
16 --positional-encoding-keys="X,Y"
17 --positional-encoding-frequencies=4
18 --mlp-width=6
19 --wavelength-key='Wavelength'
20 "X,Y,Wavelength,dHKL"
21 )
22
23 # Run careless
24 careless poly \
25   ${BASE_ARGS[@]} \
26   timepoint_{00,03,05,10,15,20,30}.s.mtz \
27   djl #output filename base

```

(a) Baseline merging script

```

- careless poly \
+ careless mono \

```

(b) Disabling harmonic deconvolution

```

- --double-wilson-parents=None,0,0,0,0,0
- --double-wilson-r=0,$R,$R,$R,$R,$R,$R

```

(c) Disabling the multivariate Wilson prior

```

- --studentt-likelihood-dof=32

```

(d) Disabling the robust error model

```

- --image-layers=2

```

(e) Disabling image layers

```

- --positional-encoding-keys="X,Y"
- --positional-encoding-frequencies=4

```

(f) Disabling positional encoding

```

- "X,Y,Wavelength,dHKL"
+ "X,Y,dHKL"

```

(g) Disabling wavelength normalization

FIG. 1. *careless* commands used in this work.

mixer.³⁸ Briefly, the DJ-1 microcrystals first pass through a microfluidic mixer where the substrate, methylglyoxal, is rapidly diffused into the crystals for reaction initiation. The freshly mixed crystals continue to flow into the sample cell observation region for data collection. The

flow speeds were adjusted to match the 3.6 μ s exposure time at a 10 Hz repetition rate so that fresh crystals intersect the x-ray interaction region for every frame. Flow rates and the position of the x-ray beam were adjusted to get different time points within the same mixer.

TABLE I. Ablation study results. To ensure consistent comparisons, all time points were truncated at 1.77 Å resolution. The average difference peak height and standard deviation across time points (after mixing) for each merge condition are shown. The average and standard deviation across time points (after mixing) for RSCC for the ligand alone and for the full model for each merge condition are also shown. The R-factors and CC_{work}/CC_{free} are reported for the 0-s time point. CC_{half} and CC_{pred} are reported over all time points.

Dataset	Active site diff peak mean \pm std. dev.	RSCC ligand only mean \pm std. dev.	RSCC full model mean \pm std. dev.	R_{work}/R_{free}	CC_{work}/CC_{free} (highest shell)
Baseline	19.0 \pm 2.6	0.960 \pm 0.005	0.873 \pm 0.014	21.94/23.58	0.510/0.449
No deconvolution	20.1 \pm 2.2	0.957 \pm 0.006	0.856 \pm 0.004	21.67/23.56	0.535/0.491
No multivariate prior	11.6 \pm 2.0	0.952 \pm 0.010	0.862 \pm 0.012	22.83/24.98	0.335/0.357
No robust error model	13.7 \pm 1.8	0.941 \pm 0.009	0.841 \pm 0.013	23.23/26.00	0.469/0.447
No image layers	14.4 \pm 2.1	0.962 \pm 0.007	0.881 \pm 0.017	21.29/22.97	0.526/0.435
No positional encoding	17.9 \pm 1.7	0.933 \pm 0.012	0.815 \pm 0.014	25.30/28.27	0.473/0.480
No wavelength norm.	9.2 \pm 1.8	0.955 \pm 0.006	0.861 \pm 0.016	22.04/24.18	0.493/0.446
Dataset	CC_{half} (highest shell)	CC_{pred} (Pearson) test/train (overall)	CC_{pred} (Spearman) test/train (overall)		
Baseline	0.347	0.723/0.685	0.548/0.554		
No deconvolution	0.315	0.744/0.706	0.553/0.559		
No multivariate prior	0.094	0.486/0.834	0.545/0.563		
No robust error model	0.305	0.766/0.977	0.523/0.533		
No image layers	0.312	0.882/0.889	0.537/0.542		
No positional encoding	0.341	0.724/0.686	0.536/0.545		
No wavelength norm.	0.340	0.703/0.750	0.518/0.531		

Diffraction images were collected at 3, 5, 10, 15, 20, and 30 s post mixing, plus an initial state (0 s, without mixing), yielding 7 datasets total. As previously reported,³⁷ DJ-1 crystallizes in space group $P3_121$, which has an indexing ambiguity.

B. Data reduction

Data were initially processed using BioCARS' Python script¹³ for hit finding as well as indexing and integration of images in parallel with Precognition software (Renz Research, Inc.). Next, indexing ambiguities were resolved for each dataset using a custom program (HEX-AMBI, M. Schmidt, personal communication). All details about processing with `careless` can be found on Zenodo (10.5281/zenodo.10481982). Briefly, the *.ii files output from Precognition were converted to .mtz format using `to_mtz.sh`. Then, a custom program, Scramble, was used to correct the consistency of the indexing convention across all datasets (the source code for Scramble is available on GitHub, <https://github.com/Hekstra-Lab/scramble/>). In this work, we used commit 746597266febb2e8dcf1a0728abe77a47e804bc8 for Scramble and commit 389b3e8084cb5a443a90a3481d80d4197b5b02b3 for `careless`. In addition to being freely available from GitHub, the source code for both of these programs is archived in our Zenodo deposition. `Careless`³³ was used to merge and scale the data.

C. Ablation study and hyperparameter sweeps

In order to assess the impact of our `careless` hyperparameter choices on performance, we performed an ablation study wherein we disabled one hyperparameter at a time. The ablations were achieved by changing the baseline merging script as detailed in Fig. 1. Hyperparameter sweeps were performed for the multivariate prior and for the robust error model. For the multivariate prior, the r value in line 3 of Fig. 1(a) is varied with the following values: 0.500, 0.800, 0.900, 0.950, 0.990, and 0.999. For the robust error model, the degrees of freedom are varied with the following values: 4, 8, 16, 32, and 64 [Fig. 1(d)]. Dataset quality was assessed with the active site difference peak height, real-space correlation coefficients (RSCCs) over the ligand and the whole model, R-factors, CC_{half} , CC_{work}/CC_{free} , and CC_{pred} . To facilitate comparison, all the experiments conducted here were truncated at a consistent, 1.77 Å resolution cutoff. These results are presented in Table I. The active site peak heights were calculated from difference maps between each time point post mixing with the initial, 0-s dataset. The average difference peak height and standard deviation across all time points for each merge condition is shown in Table I. The RSCC values were calculated against the refined models for each time point (after mixing) using the baseline script [Fig. 1(a)]. The average and standard deviation across these time points are reported. The R-factors and CC_{work}/CC_{free} are reported for the 0-s time point. CC_{half} and CC_{pred} are reported over all time points.

D. Calculation and quantification of time-resolved difference maps

We estimated phases for difference map peaks by refinement of a ground state model³⁸ against the 0-s time point data. We restricted refinement to rigid body and individual, isotropic atomic displacement parameters in the PHENIX³⁹ software package. We used the `rs.diffmap` function available in the `rs-booster` package for the `reciprocalspaceship`⁴⁰ Python library. These difference maps use the

model phases from PHENIX and the merging output from `careless`. We quantified difference map peaks using `rs.find_peaks` which relies on GEMMI.⁴¹ In Table I, we report the standard deviation of difference-peak heights across time points.

III. CARELESS ABLATION STUDIES

Modern machine-learning algorithms are complex and contain many settings that control their performance. These “hyperparameters” are fixed during model training and so must be selected or “tuned” by iteratively re-fitting a model with different settings to achieve the best results. With many settings to choose, it is generally impossible to explore the entirety of hyperparameter space for a specific problem. The reality is that determining a successful protocol for training a deep-learning model often relies on a mixture of intuition and one-dimensional hyperparameter searches.

Here, we present a protocol for merging time-resolved DJ-1 data based on one-dimensional hyperparameter searches and our developed intuition. In order to rationalize the impact of each decision involved in selecting our protocol, we conducted an ablation study wherein we disable one model feature at a time relative to the optimized protocol which we will refer to as “baseline.” This method is frequently employed in the deep-learning literature. A famous example of this principle is summarized in Fig. 4(a) of Jumper *et al.*⁴² The parameters used in our baseline protocol are listed in Fig. 1(a), which uses the same syntax as the `careless` command line interface.

The “best” parameters for merging with `careless` remains an active field of research and may be dataset-specific. Both the ablation study and hyperparameter sweep allowed us to assess the impact of each hyperparameter on the success of merging. We hope these results help future users build an intuition for the most critical parameters, demonstrate a sensible way to approach screening parameters, and share practical advice on how to evaluate the results.

We used the following metrics to judge merged dataset quality: mean and standard deviation of active site difference peak across time points (after mixing), mean and standard deviation of the real space correlation coefficient (RSCC) across time points (after mixing) for the ligand only and for the full model, R_{work}/R_{free} (for the 0-s time point), CC_{work}/CC_{free} (for the 0-s time point), CC_{half} (across all time points), and CC_{pred} (across all time points). Apart from CC_{pred} these parameters are standard crystallographic figures of merit.

In `careless`, no prior is placed on either the parameters or the output of the neural network, which estimates reflection scales. This is a design choice intended to encourage the neural network to overfit before the structure factors do. The goal is to prevent the model from introducing artifacts into the output. In consequence, in the overfit regime, the structure factors will not contain the maximum possible signal from the sample. Careful assessment of overfitting is therefore warranted to recover optimal time-resolved differences. CC_{pred} can be used to assess overfitting by evaluating the correlation between predicted and observed reflections,³³ which is an important aspect of hyperparameter selection in any machine-learning model. More specifically for CC_{pred} we report two approaches, Pearson's correlation with inverse variance weights [$CC_{pred}(Pearson)$] and Spearman's rank correlation [$CC_{pred}(Spearman)$]. We report these coefficients based on the training data and a set of test reflections, which were used to train the scaling model. Worse performance on the test data is suggestive of overfitting, which is an indicator of suboptimal hyperparameter values⁴³ (Chapter 7). All of our results are summarized in Table I, for

details on calculating each metric, see the Methods section. The outcome of each ablation study and, when applicable, one-dimensional parameter sweeps are presented below in more detail.

A. Harmonic deconvolution

One crucial consideration with polychromatic diffraction is that reflection observations need not correspond to exactly one Miller index. In fact, the pink-beam geometry maps all the reflections from a particular central ray onto the same point on the detector. A central ray is defined as the set of reflections, which lie on a ray extending outward from the origin of reciprocal space. For instance, the three reflections indicated in Fig. 2 all share the central ray passing through reflection $h = 0, k = 1$. These three reflections are called harmonics.^{44,45} In the experiment, they will be stimulated by different wavelengths of x rays. This can be seen by noting the difference in the magnitude of their scattered beam wavevectors represented by the length of the arrows in Fig. 2. The corresponding wavelengths are tabulated in the table to the right of the figure. Despite the difference in the magnitudes of the scattered wavevectors, the direction is shared. That is, the colored lines in Fig. 2 are parallel. In reality, there is only one crystal to serve as the origin for these vectors. Therefore, the reflections end up precisely superposed on the detector. This phenomenon of harmonic overlap is distinct from the spatial overlaps, which occur frequently in pink-beam crystallography.^{45,46} Spatial overlaps can be resolved during integration by profile-fitting algorithms. However, harmonic overlaps must be deconvolved during scaling.⁴⁷ *Careless*³³ implements one such algorithm for harmonic deconvolution. In order to assess the impact of harmonic deconvolution on the analysis of our time-resolved data, we ran *careless* in monochromatic mode using the *careless* mono subprogram in place of the *careless* poly subprogram typically applied to pink-beam data [Fig. 1(b)]. We found that disabling harmonic deconvolution did not markedly affect our results. Interestingly, there is a small decrease in RSCC Full Model, but otherwise, we saw no significant difference in the active site difference peak height, RSCC Ligand Only, or refinement R-values (Table I). Harmonic deconvolution did lead to a slight improvement in CC_{half} in the highest resolution bin. We attribute the lack of a dramatic difference to the high redundancy of these data as well as the experimental design. Specifically, the current beam parameters at BioCARS¹² mean that only a very small fraction (1% or less) of observations are typically harmonics. In general, we still recommend using harmonic deconvolution for pink-beam diffraction. Nevertheless, it will most likely only provide a substantial improvement for low-redundancy data or higher spectral-bandwidth sources.

B. Multivariate prior

The Bayesian model implemented in *careless* allows the user to express various expectations about the relation between crystallographic datasets through a multivariate prior distribution. To do so, users may specify a conditional dependence structure wherein each dataset is dependent on at most one “parent” dataset. Conditional dependence is an appropriate assumption in situations whenever a structure is solved simultaneously with a derivative structure. For each parent-child pair, users specify a value between zero and one for the multivariate prior r parameter, which indicates the expected degree of correlation between the two datasets (this distribution is referred to as

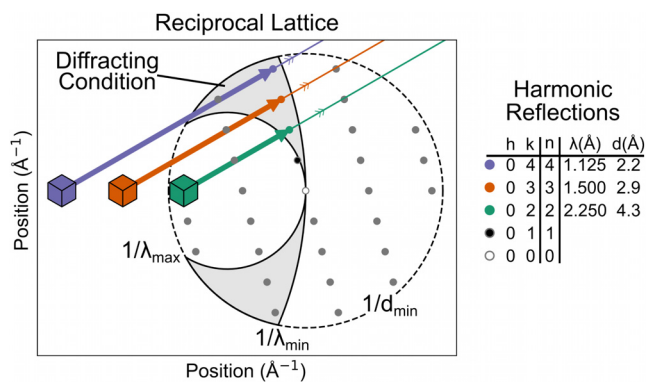


FIG. 2. The origin of harmonic overlap in pink-beam crystallography depicted by the reciprocal lattice for a 2-dimensional crystal. The diffracting condition, shaded gray, is bounded by the three limiting spheres determined by the minimum and maximum wavelength of the spectrum and the resolution limit of the sample. Three reflections from the same central ray are on the diffracting condition. Their scattered beam wavevectors, depicted as arrows, have different magnitudes, inverse wavelengths, but are all parallel. The wavelengths and resolutions of these reflections are recorded in the table to the right. Because the scattered beams are parallel, they will arrive at the same location on the detector.

the “double-Wilson” distribution, as it extends the conventional Wilson distribution to the bi- and multivariate case). On the command line, this is done by defining the parent-child relationship with `-double-wilson-parents=None, 0, 0, 0, 0, 0, 0`, where each entry in the list indicates the index of the parent dataset (datasets are indexed in the order in which they are provided to *careless*, starting at 0). In this case, each dataset is a child of the 0th dataset, which corresponds to the 0-s dataset before substrate addition. None indicates that the 0th dataset lacks a parent. Next, the r parameter is set by `-double-wilson-r=0, 0.99, 0.99, 0.99, 0.99, 0.99, 0.99`, where for each child node, the expected correlation with its parent dataset node is set {here 0.990 for all nodes except the 0th [Fig. 1(c)]}. Higher numbers indicate a greater expectation of correlation. The multivariate prior r is related but not identical to the expected Pearson correlation. As with other hyperparameters discussed here, r should be selected on the basis of crossvalidation by assessing the model fit to the training data and a held-out set of test data using the CC_{pred} measure. The multivariate prior distribution will be described in greater detail in a separate manuscript.⁴⁸

Removing the multivariate prior had a drastic impact on the quality of our analysis. The active site difference peak mean had a 7.4σ reduction. There is, however, only a minor decrease in the RSCCs. Notably, the multivariate prior was the only hyperparameter to have a considerable effect on CC_{half} , and thus, also impacts the overall resolution of the final dataset as CC_{half} was used as the main criterion to determine the resolution cutoff.

In addition to the ablation study, we also performed a one-dimensional parameter sweep of the prior correlation value (Fig. 3). We used the height of the active site difference peak as our main criterion for selecting the optimal value. We found a maximum in difference peak height with a prior correlation value of 0.990 [Fig. 3(a)]. At this value, the mean value across all time points was 19σ , whereas the maximum value was 23σ (15 s time point). Additionally, we far surpass the noise standard of 3σ at our optimized value (0.990) and meet

or exceed this for all of our tested values. The active site difference peak heights, however, were strongly influenced by the prior correlation value [Table 1 and Fig. 3(a)]. We encourage users to carefully screen the prior correlation value for their own datasets to try to find a maximum.

As an additional real space metric, we used RSCCs for the ligand alone and for the full model [Figs. 3(b) and (c)]. For the ligand alone, the RSCC increases monotonically with respect to r , but the RSCC for the full model starts to suffer around $r = 0.950$. We also used R_{free} as an additional metric to assess map-model agreement and see the lowest values between 0.950 and 0.999.

$CC_{pred}(Pearson)$ and $CC_{pred}(Spearman)$ were both computed for crossvalidation to assess overfitting and to ultimately decide the best r value. Ideally, these CC_{pred} indicators should have the highest possible value while having the smallest gap between the test and train set, as this would indicate that the predicted and observed reflections correlate well with each other. In order to estimate the uncertainty of our correlations, we used bootstrapping⁴⁹ to estimate a distribution for both $CC_{pred}(Pearson)$ and $CC_{pred}(Spearman)$. For $CC_{pred}(Pearson)$, we see the smallest gaps between the test and train set median values [Fig. 3(e)] at $r = 0.990$, which corresponds well with our difference peak maximum and reasonably matches our RSCC and R_{free} results. Interestingly, r -values below 0.990 demonstrate strongly bimodal $CC_{pred}(Pearson)$ distributions, which resolve at higher values. At 0.999, the test and train values increase, but the gap between test and train also increases. We interpret this as a less favorable setting. However, we note that the discrepancy in difference peak heights between $r = 0.990$ and 0.999 is marginal.

C. Robust error model

A key consideration in fitting models to data is specifying an error model, more specifically how the measurement errors are expected to be distributed. The default choice in *careless* is that errors follow a normal distribution with a standard deviation determined by the empirical uncertainty estimates from integration. The major drawback of a normally distributed error model is that it is very sensitive to outliers (see Chapter 2.4 of Murphy⁵⁰). Fortunately, the Bayesian framework used by *careless* is quite flexible and affords some control over the error model. Specifically, *careless*³³ provides a robust alternative to the normally distributed error model in the form of a Student's t -distribution. The Student's t -distribution has heavier tails than a normal and is therefore more tolerant of outliers.^{51,52} The degree of tolerance can be controlled by specifying a parameter of the t -distribution, the degrees of freedom, ν . For larger values of degrees of freedom, the model is more sensitive to outliers. As ν approaches infinity, the error model becomes equivalent to the normal distribution. In nearly all examples we have encountered, $\nu = 32$ has outperformed the normally distributed error model. Therefore, we recommend this setting, which can be implemented on the command line using `-studentt-likelihood-dof = 32` [Fig. 1(d)]. In our ablation study, we found that the robust, Student's t -distributed error model was among the strongest determinants of model performance as judged by time-resolved difference peak height (Table 1). Specifically, the use of the normal distribution in place of the robust error model resulted in a 5.3σ decrease in the mean difference peak height. There was also a minor decrease on the RSCC Ligand Only and RSCC Full Model.

We performed a one-dimensional parameter sweep of ν to determine whether 32 was the optimal value for our data (Fig. 4). This was implemented by simply changing `-studentt-likelihood-dof =` to either 4, 8, 16, or 64. We used the active site difference peak heights, RSCCs, R_{free} , and both $CC_{pred}(Pearson)$ and $CC_{pred}(Spearman)$ as the main metrics to assess the results (Fig. 4). We again used bootstrapping⁴⁹ to estimate a distribution for both $CC_{pred}(Pearson)$ and $CC_{pred}(Spearman)$. By all these criteria, it is immediately clear that Student's t distribution outperforms the normal distribution, just as demonstrated in Dalton *et al.*³³ Upon closer inspection, it is evident that 32 is the best value for the degrees of freedom of this dataset. Although the active site difference peak heights are not strongly affected by ν , there is a moderate maximum at 32 [Fig. 4(a)]. The RSCCs are more sensitive to ν and both show a moderate peak at 32 [Figs. 4(b) and 4(c)]. The average R_{free} value across time points at 64 is slightly lower than at 32 [Fig. 4(d)]. For $CC_{pred}(Pearson)$, 32 has the smallest difference between the mean test and train and has the train set with the least variability [Fig. 4(e)]. For $CC_{pred}(Spearman)$, there is a slight maximum in the test and train mean values and the smallest gap between the test and train sets at 32 [Fig. 4(f)]. Overall, this demonstrates that utilizing a few figures of merit is a good practice for selecting values from one-dimensional parameter sweeps. We particularly recommend this combination of using a real-space measure like difference-peak height alongside CC_{pred} as criteria.

D. Image layers

By default, *careless* uses a scaling model with purely global parameters. This can be inappropriate for serial crystallography applications wherein each diffraction image typically corresponds to a separate crystal. Variation in the size and quality of samples require different scaling corrections. In such situations, it is important to allocate some of the parameter budget to local parameters, which are able to make corrections to each sample independently. In *careless*, this can be done by using image layers. Image layers are neural network layers, which have separate parameters for each image in the dataset [see Fig. 5(a) of Ref. 33]. The appropriate number of image layers for a dataset can be determined by crossvalidation. In this study, we used 2 image layers which we have found to be sufficient in most serial-crystallography cases. This was implemented by using the `-image-layers = 2` command line argument [Fig. 1(e)]. Crucially, we have never seen 2 image layers lead to overfitting for a serial dataset. In the single-crystal scenario, we have observed cases where image layers can be detrimental to difference map inference. Regarding the serial DJ-1 data, removing the image layers led to a 4.6σ decrease in the signal-to-noise of the time-resolved difference peak supporting our assertion that local parameters are beneficial for serial-crystallography data analysis (Table 1). There is, however, no marked change on the RSCCs or other crystallographic statistics.

E. Positional encoding

Owing to a variety of effects, the observed intensity of a reflection can vary spatially across the detector.³¹ In the case of features such as panel gaps, background scatter, and shadows from the beamstop or other equipment, intensity variations can be abrupt. These sorts of high-frequency variations can be difficult for neural networks to

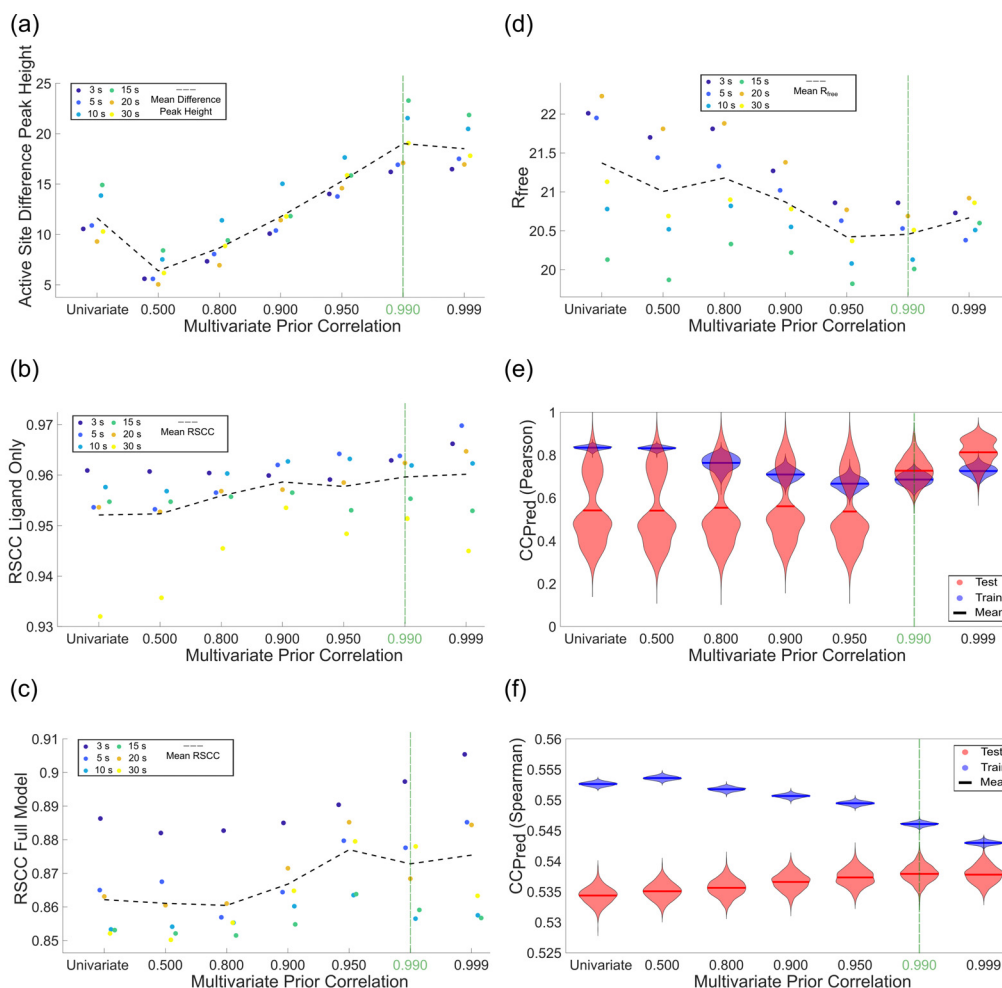


FIG. 3. Results of a one-dimensional sweep of the multivariate prior correlation parameter. (a) The active site difference peak heights for various multivariate prior correlation values. There is a clear maximum at 0.990. (b) and (c) RSCCs for ligand alone and for the full model. (d) R_{free} is provided as an additional global measure of map to model agreement. (e) and (f) CC_{Pred} was calculated to assess overfitting using either the maximum-likelihood weighted Pearson or Spearman correlation coefficients. Careless produces a posterior distribution for each intensity observation. The mean and standard deviation of this distribution are recorded in the *_predictions_#.mtz files saved after model training. The means of these distributions is typically used to compute CC_{Pred} or the correlation between observed and predicted intensities. Here, we quantify the uncertainty in CC_{Pred} using the bootstrap method whereby we resample the predicted intensities recorded in the careless output with replacement 1000 times yielding 1000 estimates of CC_{Pred} per hyperparameter setting. These bootstrapped estimates are visualized as violin plots. The optimal hyperparameter setting has the highest value while having the smallest gap between test and train. We observe this clearly at 0.990 for $CC_{Pred}(Pearson)$ and at 0.999 for $CC_{Pred}(Spearman)$, indicating that the exact optimum likely lies between the two.

model.⁵³ One well-validated solution to this problem is to use a positional encoding strategy^{54–56} whereby coordinates are mapped to a higher-dimensional representation which the scale model can more easily interpret. In *careless*, this is accomplished using the `-positional-encoding-keys=` command line argument. This parameter takes a set of comma separated metadata keys and instructs *careless* to encode them into a higher-dimensional representation [Fig. 1(f)]. The dimensionality of the encoding is controlled by the `-positional-encoding-frequencies=` argument [Fig. 1(f)]. In our most successful protocol, we used positional encoding of the detector coordinates of each reflection observation. We found this gave a modest, approximately 2σ , increase in difference peak height (Table I). Interestingly, removing positional encoding negatively affected RSCCs,

R_{work}/R_{free} and CC_{work}/CC_{free} , so its main benefit seems to contribute to absolute scaling accuracy rather than the accuracy of structure factor differences.

F. Wavelength normalization

At synchrotron light sources, it is often necessary to use a polychromatic x-ray beam in order to record time-resolved diffraction at sufficient temporal resolution for biological processes. The BioCARS 14-ID beamline at the Advanced Photon Source provides a polychromatic beam with sub-microsecond time-resolution (best time resolution at this beamline is determined by a duration of the single x-ray pulse).^{12,13} Owing to the nature of synchrotron radiation, the

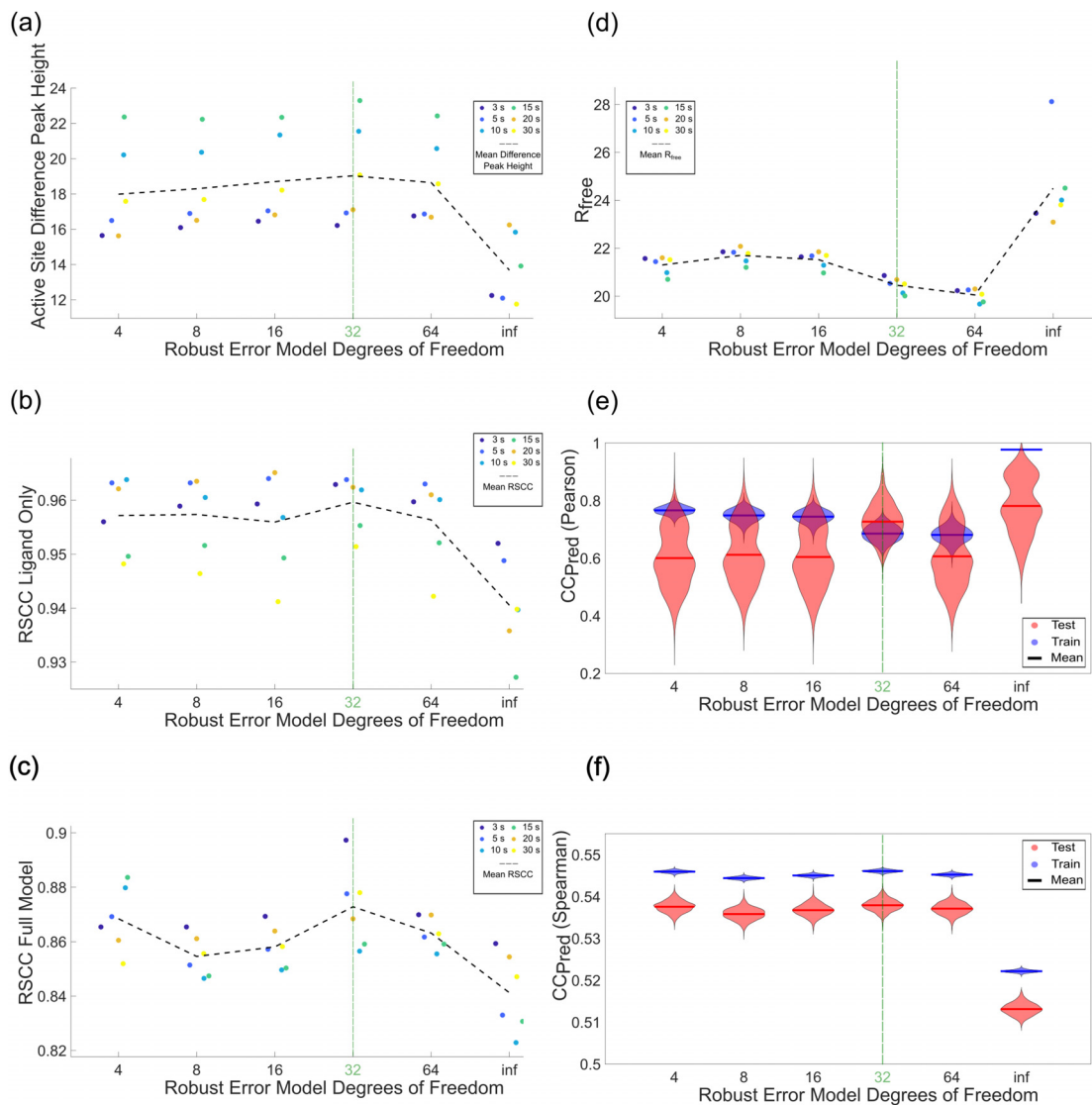


FIG. 4. The results of the one-dimensional sweep of Student's t degrees of freedom, ν , on active site difference peak heights, RSCCs, and CC_{Pred} . (a)–(c) The active site difference peak heights and RSCCs have a slight maximum at 32. (d) R_{free} is provided as an additional global measure of map to model agreement. (e) $CC_{\text{Pred}}(\text{Pearson})$ has the smallest gap between test and train at 32, which are the optimal results. (f) $CC_{\text{Pred}}(\text{Spearman})$ is an alternative metric to assess overfitting. It has the best results at 32 with the highest overall value, albeit by a small margin, and the smallest gap between test and train. The distributions visualized in the violin plots were generated by the bootstrap method described in the caption of Fig. 3.

BioCARS beam exhibits a characteristically skewed undulator spectrum [Fig. 5(a)]. The relatively large bandwidth of the beam provides significantly increased photon flux compared to monochromatic beamlines and consequently much shorter exposure times. It also increases the number of Bragg peaks observed in each image [Fig. 5(b)]. However, the spectrum bandwidth introduces an additional challenge in data analysis. In particular, the basal flux differs by wavelength. Therefore, the empirical intensity of each reflection depends on the wavelength at which it is observed on a particular image. In order to model this, a straightforward approach can be used in *careless*. Namely, the empirical wavelength of each

reflection observation is computed from the experimental geometry used at integration. Then, this wavelength is provided to the *careless* scale model during merging. On the command line, this is done by including 'Wavelength' as a metadatum [Fig. 1(g)]. Our experiments indicate that *careless* is able to approximate the spectral nature of the beam. This can be best visualized as a histogram of the scale values estimated by *careless* as a function of wavelength [Fig. 5(c)]. As can be seen, the inferred scales recapitulate the characteristic peak and long tail of the pink-beam spectrum. Removing the wavelength metadatum from the *careless* scaling model leads to a featureless dependence of scale on

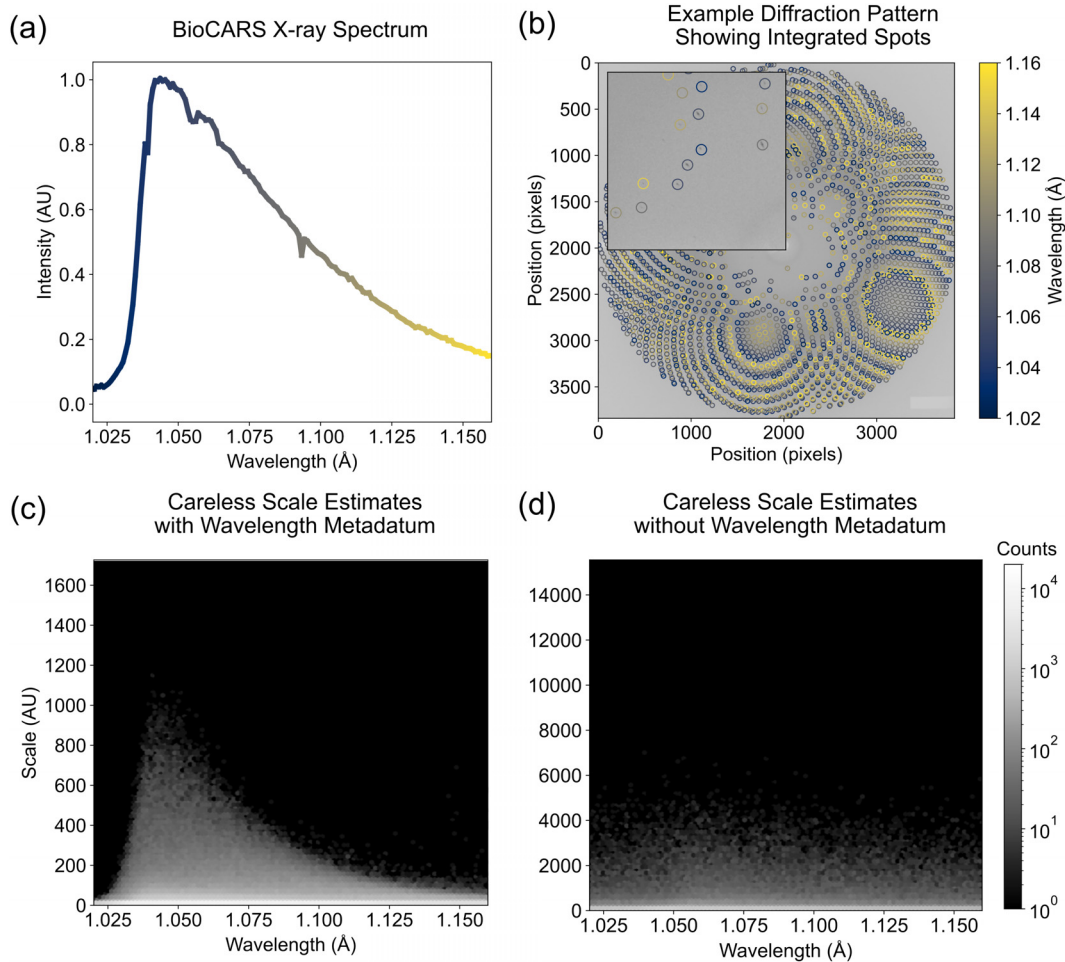


FIG. 5. *Careless* infers the spectrum from wavelength metadata. (a) The spectrum of the x-ray beam at BioCARS measured by a channel-cut monochromator. (b) An example diffraction pattern from our DJ-1 dataset with the indicated spots circled and colored by the predicted peak wavelength. (c) A two-dimensional histogram of the scale value (systematic error) predicted by *careless* and the peak wavelength predicted by Precognition (Renz Research). The skewed distribution is similar to the BioCARS spectrum, indicating that *careless* can infer spectral information. (d) The equivalent two-dimensional histogram produced from the wavelength-normalization ablation study wherein *careless* did not have access to the peak wavelength of each reflection observation.

wavelength [Fig. 5(d)], indicating that *careless* is not able to latently infer the wavelength from the remaining metadata. Furthermore, removing the wavelength leads to the worst performing model in our ablation study with time-resolved difference peaks diminished by roughly 10σ relative to our baseline. There is, however, no major change on RSCCs or other crystallographic metrics.

IV. DISCUSSION

We demonstrated the importance of various *careless* hyperparameters for getting high-quality merged datasets and especially for capturing time-resolved differences. In particular, we found that wavelength normalization had the largest impact on difference peak heights. This is due to our use of a wide bandwidth ($\sim 5\%$) x-ray beam. The next most important parameter, as judged from difference peak heights, is the use of a multivariate prior. For our dataset, we found a

prior correlation of $r = 0.990$ to be the optimal value, but in general, we found the specific value to be quite influential. It is possible that the ideal value may vary for different datasets and types of time-resolved experiment modalities. It is strongly recommended to do a one-dimensional sweep of the multivariate prior to determine which value is best for your specific use case. CC_{pred} -based crossvalidation measures can be unreliable for the multivariate prior. In the limit $r \rightarrow 1$, the structure factor differences between time points are forced to zero. This equates to learning a single set of structure factors to represent the average across all the time points. Therefore, higher r -values can be thought of as reducing the effective parameter count of the model, which leads to less apparent overfitting and superior CC_{pred} values. This indicates a shortcoming of CC_{pred} in determining the optimal r -value. We caution users that for this scenario, CC_{pred} should not be trusted uncritically. If real-space measures of performance are available,

such as the difference peak heights, this is one place where it would be appropriate to rely on those rather than assessments of model fit.

Using a robust error model (Student's t likelihood with $\nu = 32$) and implementing image layers (with a value of 2) are additional settings, which may improve difference map peak heights. These specific values have been consistent for the serial crystallography datasets to which `careless` has been applied thus far,³³ so they are a good starting point. Given the consistency in the performance of these values, one-dimensional sweeps are not necessarily vital for these hyperparameters, especially when using the default parameters for the neural network depth and width (`-mlp-width=10` and `-mlp-layers=20`). Our `mlp-width` was kept at 6 due to limits on our accelerator card and the large number of reflections across the seven datasets. We recommend you increase `mlp-width` to the maximum value supported by your accelerator card.

Positional encoding had only a modest impact on the difference peak heights, but it had a larger impact on standard crystallographic figures of merit, like R_{work}/R_{free} and CC_{work}/CC_{free} . Thus, it is still an important parameter to utilize. Harmonic deconvolution had minimal effects on the analysis quality. This is expected on the basis of the experimental design in this case, in which only a small fraction ($<1\%$) of reflections have contributions from multiple harmonics.

RSCCs can be used as an additional real-space metric. Overall, the RSCCs provide similar, but not identical, results as the active site difference peak heights. Specifically, we see slight differences in the results for the one-dimensional hyperparameter sweep for the multivariate prior and the ablation studies for image layers and wavelength normalization. Given the resolution and the small scale of change expected for these data, the active site difference peak height is a more appropriate measurement. For other experiments, especially those with larger structural changes, the RSCC Full Model may be a better metric. We recommend calculating the active site difference peak heights and the RSCCs and then using experimental knowledge, R_{pred} and CC_{pred} to determine optimal hyperparameter settings, especially for the multivariate prior. For example, in our one-dimensional hyperparameter sweep of the multivariate prior, the average difference peak height has a peak at 0.990, the average RSCC for the ligand has a maximum at 0.999, and the average RSCC for the full model has a slight peak at 0.950. When being conservative about overfitting and utilizing CC_{pred} , it becomes clear that 0.990 is an appropriate choice.

The resolution limit, `dmin`, of the data also needs to be set for every iteration of `careless`. We found that the maximum resolution attainable, based on a CC_{half} threshold of 0.3, was dependent on the hyperparameter values we used. Our baseline dataset has 1.77 Å resolution across all time points. We also processed our data with CrystFEL.^{57,58} The resolution cutoff for each time point in this case was determined based on >0.2 value of CC_{half} and had a slightly different value for each time point, but the average resolution was 1.98 Å. We attribute this 0.21 Å gain in resolution to the structure imposed by the multivariate prior. This is clearly demonstrated by the degradation of highest-shell CC_{half} exhibited in the multivariate prior ablation (Table I). Since it is ideal to have as high resolution data as possible to interpret time-resolved differences, we see this as a strong argument in favor of using `careless` for time-resolved crystallography data.

Based on our experience, we recommend the following protocol. Starting from an appropriate baseline configuration, such as the one presented in Fig. 1(a),

- (1) determine the resolution cutoff using CC_{half} ;
- (2) sweep the multivariate prior correlation (r) and determine the optimal value based on real-space performance measures;
- (3) determine the optimal value of ν , Student's t likelihood degrees of freedom, based on CC_{pred} ; and
- (4) re-determine the resolution cutoff using CC_{half} .

For the first resolution cutoff test (1), we suggest starting with the resolution determined from a standard program, such as CrystFEL,⁵⁷ Precognition (Renz Research, Inc.), DIALS,⁵⁹ or the Daresbury Laue Suite,⁶⁰ and then increase this in small increments determined by your desired precision until you approach a CC_{half} value of 0.3 in the highest resolution shell. For the one-dimensional sweep of the multivariate prior correlation parameter (2), we recommend the following values: 0.500, 0.800, 0.900, 0.950, 0.990, and 0.999. The optimal value should be chosen to avoid significant overfitting as judged by CC_{pred} while maximizing a desirable real-space measure such as difference peak height or RSCC.

Once a suitable multivariate Wilson parameter has been established, Student's t distribution ν sweep can be performed (3). It is most reasonable to test logarithmically spaced values of ν . For instance, $\nu = 4, 8, 16, 32, 64$. CC_{pred} seems to be the most useful metric for selecting the optimal ν value. Finally, once these hyperparameters have been optimized for your dataset, it may be possible to gain additional resolution. It may be worthwhile to attempt to extend the resolution (4) in small increments (0.01–0.05 Å) until CC_{half} in the highest resolution shell drops below 0.3.

While performing the aforementioned hyperparameter sweeps, we suggest calculating all the figures of merit in Table I to assess your results. Sometimes, the various figures of merit can give conflicting results for the overall best dataset. For example, sometimes, the R -factors would decrease (improved result), but the difference peak heights would also decrease (worse result). Whenever they are available, we suggest using difference peak heights or RSCCs as the criterion for selecting the best strategy. We view CC_{half} and CC_{pred} as the next two most important metrics. CC_{half} acts as a standard crystallographic metric to help evaluate the overall consistency and resolution of the dataset.

CC_{pred} is useful to assess overfitting and can be estimated using either Pearson's or Spearman's rank correlation coefficient. Which estimator is more accurate is likely to depend on the particular data being analyzed. Pearson's method admits weights derived from the empirical uncertainties of reflection intensities estimated during integration. Spearman's rank correlation coefficient is generally more robust to outliers. Therefore, the quality of the uncertainty estimates and the frequency of outliers in a dataset will impact the relative accuracy of each metric. More study is needed to assess alternative measures of fit. However, for the time being, we recommend users consider both Pearson and Spearman CC_{pred} for hyperparameter optimization. Ultimately, a decision can frequently be made as to which is superior on the basis of real-space measures.

Overall, we have presented insights on the key parameters when utilizing `careless` to merge time-resolved serial crystallography data. We have also provided practical advice on how to screen such parameters and how to interpret the results. Although our datasets were collected via pink-beam crystallography at a synchrotron, we expect the described approach to be broadly applicable across serial crystallography modalities. Additionally, we believe there is a general takeaway for scaling and merging time-resolved data regardless of the

algorithm being used. In particular, the importance of the multivariate prior for these datasets demonstrates that scaling all time points jointly, and ideally against the same reference data, is generally useful for time-resolved crystallography data. We expect this conclusion applies equally to data scaled by variational inference or more traditional, least squares methods.

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AUTHOR DECLARATIONS

Conflict of Interest

The authors have no conflicts to disclose.

Author Contributions

Kara A. Zielinski: Data curation (equal); Investigation (equal); Methodology (equal); Software (equal); Validation (equal); Visualization (equal); Writing – original draft (equal); Writing – review & editing (equal). **Cole Dolamore:** Data curation (equal); Investigation (equal); Writing – review & editing (equal). **Harrison K. Wang:** Validation (equal); Visualization (equal); Writing – review & editing (equal). **Robert W. Henning:** Data curation (equal); Funding acquisition (equal); Resources (equal); Writing – review & editing (equal). **Mark A. Wilson:** Funding acquisition (equal); Resources (equal); Writing – review & editing (equal). **Lois Pollack:** Funding acquisition (equal); Resources (equal); Writing – review & editing (equal). **Vukica Srajer:** Data curation (equal); Funding acquisition (equal); Methodology (equal); Resources (equal); Software (equal); Validation (equal); Writing – original draft (equal); Writing – review & editing (equal). **Doek R. Hekstra:** Funding acquisition (equal); Methodology (equal); Resources (equal); Software (equal); Validation (equal); Visualization (equal); Writing – original draft (equal); Writing – review & editing (equal). **Kevin M. Dalton:** Conceptualization (lead); Data curation (equal); Funding acquisition (equal); Investigation (equal); Methodology (lead); Project administration (lead); Resources (equal); Software (lead); Supervision (lead); Validation (equal); Visualization (equal); Writing – original draft (equal); Writing – review & editing (equal).

DATA AVAILABILITY

The data that support the findings of this study are openly available in Ref. 61.

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