

RESEARCH

Open Access



Structural properties of generalised Planck distributions

Anthony G. Pakes

Correspondence:

tony.pakes@uwa.edu.au

Department of Mathematics & Statistics, Univ. of Western Australia, 35 Stirling Highway, Crawley 6009, WA, Australia

Abstract

A family of generalised Planck (GP) laws is defined and its structural properties explored. Sometimes subject to parameter restrictions, a GP law is a randomly scaled gamma law; it arises as the equilibrium law of a perturbed version of the Feller mean reverting diffusion; the density functions can be decreasing, unimodal or bimodal; it is infinitely divisible. It is argued that the GP law is not a generalised gamma convolution. Characterisations are obtained in terms of invariance under random contraction of a weighted version of a related law. The GP law is a particular instance of equilibrium laws obtained from a recursion suggested by a genetic mutation-selection balance model. Some related infinitely divisible laws are exhibited.

Keywords: Planck distribution, Mean reverting diffusion, Modal properties, Infinite divisibility and self-decomposability, Length biased and weighted laws, Characterisation, Mutation-selection balance

AMS Subject Classification: 60E05; 62E10; 60E07; 60J60; 92D10

1 Introduction

In 1900 Planck derived a formula describing the energy spectrum of black body radiation. The energy density at frequency ν of a body at temperature T (in degrees Kelvin) is

$$\rho(\nu, T) = \frac{8\pi h}{c^3} \cdot \frac{\nu^3}{e^{h\nu/kT} - 1} \quad (\nu, T > 0) \quad (1)$$

where h is Planck's constant, k is Boltzmann's constant, and c is the velocity of light. Good accounts of Planck's reasoning are given by Cropper (1970, Chap. 1), Longair (1984, Chap. 10) and, highly recommended, Weinberg (2013). Planck's first derivation combined reasoning based on electromagnetic theory and statistical mechanics, as well as empirical curve fitting to then new experimental results. In the last connection see Pais (1982, figure on p. 367).

The form (1) suggests defining a Planck family of distributions whose members have a probability density function (pdf) of the form

$$f(x) = f(x; \lambda, a) = \frac{x^{\lambda-1}}{K(e^{ax} - 1)} \quad (x > 0) \quad (2)$$

where $\lambda > 1$, $a > 0$ and K is the normalisation constant. This family was introduced to the world in Johnson and Kotz (1970, p. 273), but subsequently dismissed from service by omission from the second edition of this reference work. The family (2) appears indirectly in Kleiber and Kotz (2003) in that its reciprocal form was used by Davis (1941) to fit personal income data for the U.S.A. in 1918. Subsequent analysis by Champernowne has shown that his own three-parameter distribution yields a better fit. See Kleiber and Kotz (2003, p. 240). Some elementary computations with (2) are reported by Tomovski et al. (2012). Finally, cases of (2) where $\lambda = d + 1$, where $d = 1, 2, \dots$, arise in physics literature as describing black-body radiation in a d -dimensional universe. See Puertas-Centeno et al. (2017) and its references.

Nadarajah and Kotz (2006) have proposed the generalized Planck family $GP(\lambda; a, b, c)$ whose pdf is

$$g(x) = g(x; \lambda, a, b, c) = K^{-1} \frac{x^{\lambda-1} e^{-ax}}{1 - ce^{-bx}} \quad (x > 0) \quad (3)$$

where $a, b > 0$ and $-1 \leq c \leq 1$ with $\lambda > 0$ if $c < 1$, and $\lambda > 1$ if $c = 1$. The constant K is a normalisation factor which is evaluated at (6) below. Their motivation seems to be that the case $c = 0$ yields the gamma family, and hence that the generalised family gives a more malleable family of pdf's. We will extend the range of the parameter c to allow $-\infty < c \leq 1$. If $\lambda = 1$ then choosing $0 < c < 1$ (resp. $-1 < c < 0$) yields the Bose-Einstein distribution (resp. the Fermi-Dirac distribution); see Kittel (1969, Chapter 9) or Olver et al. (2010, #25.12(iii)). In addition, the $GP(\frac{1}{2}; a, b, c)$ family is implicitly present in a specific 'house of cards' model for mutation-selection balance used by Turelli (1984, §3) to study heritable variation. This term is used by Kingman (1978) to describe the situation where the rate of mutation at a locus is independent of the parental allele, thus collapsing the ancestral history of selection at that locus. Nadarajah and Kotz (2006) derive formulae for moments, the characteristic function and entropies, they mention shapes of the pdf, and they look briefly at maximum likelihood estimation.

Our objective in this paper is to investigate some structural properties of the GP laws and point out limits of this investigation. In §2 we present some fundamental identities, a representation as a scale mixture of gamma laws (Theorem 1), and limiting behaviour as $\lambda \rightarrow \infty$ (Theorem 2) or $b \rightarrow \infty$ (Theorem 3). In §3 we summarise for the readers convenience properties of the gamma law we aim to extend to the GP laws.

The gamma family occurs as the limiting/stationary law of the Feller diffusion process, i.e., the continuous-state branching process with immigration. In §3 we show that the GP laws arise in the same way through an additive non-linear adjustment of the drift term.

Nadarajah and Kotz (2006) list with no proofs how the shape of $g(x) := g(x; \lambda, a, b, c)$ depends on values of the parameters. Their assertions about multimodality are vague. Modal properties of g are described in detail in §5. Theorem 4 is an exact statement about when g is, or is not, strongly unimodal. Theorems 5–8 examine its modal properties for the respective cases $a = b$ and $c = 1$, $c = 1$, $0 < c < 1$, and $c < 0$. In summary, depending in a fairly complicated way on parameter combinations, g either is decreasing, unimodal, or bimodal.

A mixture representation (Theorem 9) and infinite divisibility properties are investigated in §6. If $\lambda \leq 2$ and $0 \leq c \leq 1$, then the $GP(\lambda; a, b, c)$ law is infinitely divisible (abbreviated to *infdiv*; Theorem 10), and the situation is unresolved if $\lambda > 2$ and $c \neq 0$.

In recent years the infinite divisibility of many common distributions has been proved by showing that they are generalised gamma convolutions (Nadarajah and Kotz (2006) or Steutel and van Harn (2004)). However, as we argue at the end of §6, that if $c \neq 0$, then the GP laws are not in general a generalised gamma convolution.

Proposition 1 gives sufficient conditions for self-decomposability of GP laws in the restricted case where g is decreasing. The proofs verify Bondesson’s (1987) sufficient conditions. Note that since self-decomposable laws are unimodal, the parameter combinations in §5 for bimodality preclude self-decomposability.

Many laws can be characterised as the fixed point of a transformation defined by weighting the law and then applying a random rescaling. Often the weighting operation is length, or size, biasing. More specifically for this case, let r be real. For any random variable $Y \geq 0$ with distribution function G and finite moment function $M_G(r) = E(Y^r)$, define the order- r length-bias operator \mathcal{L}_r which maps G into the distribution function $G_r(x) = \int_0^x y^r dG(y)/M_G(r)$. We extend our notational convention so that $\mathcal{L}_r Y$ denotes a random variable whose distribution function is G_r . If $r > 0$ then $\mathcal{L}_r Y \geq_{st} Y$. This property opens the possibility of characterising $L(Y)$ through a relation of the form $Y \stackrel{d}{=} V\mathcal{L}_r Y$ where $0 \leq V \leq 1$ is a random variable and the factors in the product are independent. We obtain such characterisations of GP laws in §7.

In §8 we study what amounts to a formal generalisation of Kingman’s (1978) ‘house of cards’ model for genetic mutation-selection balance. Equilibrium laws of this model generalise (3) and a representation (Theorem 17) generalises Theorem 9. Finally, in §9 we exhibit some discrete and continuous infdiv laws which are related in certain ways to the generalised Planck laws. Longer proofs of our various results comprise §10.

If a random variable X has the pdf (3) then we write $X \sim GP(\lambda; a, b, c)$, and similarly for other laws. The standard gamma law with shape parameter λ (i.e., $a = 1$ and $c = 0$ in (3)) is denoted by $\text{gamma}(\lambda)$, and we use the notation $\gamma(\lambda)$ for a random variable having this law. Note the tail equivalence $P(X > x) \sim P(\gamma(\lambda)/a > x)$ as $x \rightarrow \infty$. Similarly, the beta law with parameters α and β is denoted by $\text{beta}(\alpha, \beta)$. Independence of random variables X and Y is denoted by $X \perp Y$. Finally, $L(Y)$ denotes the law of a random variable Y , and $X \stackrel{d}{=} Y$ means $L(X) = L(Y)$.

There is some duplication of notation between sections, but no confusion will result from this.

2 Fundamental facts

The Lerch zeta function is defined by the series

$$\Phi(c, \lambda, \theta) = \sum_{n=0}^{\infty} \frac{c^n}{(\theta + n)^\lambda} \tag{4}$$

where $\theta > 0$, and the series converges for all positive λ if $|c| < 1$, for $\lambda > 1$ if $|c| = 1$ and it is conditionally convergent if $c = -1$ and $\lambda = 1$. The function $\Phi(1, \lambda, \theta)$ is usually called the Hurwitz zeta function, and the Riemann zeta function $\zeta(\lambda)$ is the still more special case $c = \theta = 1$. The function $(\theta + n)^{-\lambda}$ is a completely monotone function of θ and in fact it is the Laplace transform of $v^{\lambda-1}e^{-nv} / \Gamma(\lambda)$. Substituting the resulting integral representation of $(\theta + n)^{-\lambda}$ into (4) and summing gives the known integral representation

$$\Phi(c, \lambda, \theta) = \frac{1}{\Gamma(\lambda)} \int_0^\infty \frac{v^{\lambda-1}e^{-\theta v}}{1 - ce^{-v}} dv, \tag{5}$$

which clearly is well defined if $-\infty < c < 1$ and $\lambda > 0$, or if $c = 1$ and $\lambda > 1$, and $\theta > 0$ in either case. We will take (5) as the definition of the Lerch function (or Lerch transcendent as it is called by some). See Gradshteyn and Ryzhik (1980, 9.550,6) for these formulae and Olver et al. (2010, Chapter 25) for graphs and references.

There is a considerable literature devoted to representations and identities of the Lerch zeta as a function of λ , i.e., the integral (5) is regarded as a Mellin transform. More relevant for our purposes is the fact that (5) is a Laplace transform in the variable θ with c and λ understood as parameters.

Integrating the pdf (3) and setting $\alpha = a/b$ shows that the normalisation constant is

$$K = b^{-\lambda} \Gamma(\lambda) \Phi(c, \lambda, \alpha). \tag{6}$$

It follows that the Laplace-Stieltjes transform of $X \sim GP(\lambda; a, b, c)$ is

$$E(e^{-\theta X}) = \frac{\Phi(c, \lambda, \alpha + \theta/b)}{\Phi(c, \lambda, \alpha)}, \tag{7}$$

finite if $\theta > -a$. The moment function of X is

$$M_g(t) := E(X^t) = b^{-t} \frac{\Gamma(\lambda + t)}{\Gamma(\lambda)} \cdot \frac{\Phi(c, \lambda + t, \alpha)}{\Phi(c, \lambda, \alpha)}, \quad (t \geq 0). \tag{8}$$

Remark 1 *The parameter b can be understood as a scaling parameter, and for most of our purposes it can be set to unity. In cases where this simplification is made, we denote the family by $GP(\lambda; a, c)$ and the pdf by $g(x; \lambda, a, c)$.*

The quotient of gamma functions in (8) equals the moment function of $\gamma(\lambda)$, and so one conjectures that the quotient of Lerch zetas also is a moment function. The series definition shows this is the case if $c > 0$, thus giving one proof of the following result which extends the mixture representation of the family (2) observed by Johnson and Kotz (1970).

Theorem 1 *If $X \sim GP(\lambda; a, b, c)$, where $a, b, \lambda > 0$ and $0 < c \leq 1$, then*

$$X \stackrel{d}{=} \frac{\gamma(\lambda)}{N}$$

where $\gamma(\lambda) \perp N$,

$$P(N = a + n) = p(n) := (\Phi(c, \lambda, a))^{-1} \frac{c^n}{(a + n)^\lambda}, \quad (n = 0, 1, \dots), \tag{9}$$

and

$$v(t) := E(N^{-t}) = \frac{\Phi(c, \lambda + t, a)}{\Phi(c, \lambda, a)}. \tag{10}$$

Proof The pdf (3) can be written as

$$g(x) = \sum_{n \geq 0} P(N = a + n) \frac{x^{\lambda-1} (a + n)^\lambda}{\Gamma(\lambda)} e^{-(a+n)x},$$

a mixture of scaled gamma pdf's, that is, $X = \gamma(\lambda)/(a + n)$ with probability $p(n)$. This is the essence of (9). The moment function (10) arises by multiplying (9) by $(a + n)^{-t}$, summing, and appealing to (4). \square

Let N_L denote a random variable having the Lerch law meaning that for $n = 0, 1, \dots$, $P(N_L = n) = p(n)$ as defined in (9). See Johnson et al. (2005, pp. 526–530) and the

references there for information about the Lerch law and its special cases. Our designation ‘Lerch law’ for $L(N_L)$, where $N_L = N - a$, is adopted from Gupta et al. (2008). It is called a discrete Pareto law by others, e.g., Steutel and van Harn (2004, p. 68) for the case $\alpha = c = 1$. It is shown in these references that $L(N_L)$ is compound Poisson. The Lerch law has been shown to provide a better fit to certain biological count data than the generalised Poisson law and it is discussed as the limiting law of certain positive-recurrent birth and death processes by Klar et al. (2010). With this definition we see that the random variable N in Theorem 1 has a shifted Lerch law, $N = N_L + a$.

Suppose that X_i ($i = 1, 2$) has the $GP(\lambda_i; a_i, c_i)$ law with $0 \leq c_i \leq 1$, then it follows from Theorem 1 that $X_i \stackrel{d}{=} \gamma_i(\lambda_i)/N_i$ where N_i has a shifted Lerch law. If $X_1 \perp X_2$, then $X_1/X_2 \stackrel{d}{=} WD$ where $W = \gamma_1(\lambda_1)/\gamma_2(\lambda_2)$ and $D = N_1/N_2$. It follows that $W \perp D$, W has a type-2 beta law, and D has a discrete law which assigns positive mass to points comprising a countable dense subset of $[0, \infty)$. This outcome was noted by Johnson and Kotz (1970) for the family (2).

Theorem 1 has another interesting consequence. Let V have a standard normal law and $V \perp \gamma(\lambda)$. Then $V\sqrt{2\gamma(\lambda)}$ has the normal mixture law (Steutel and Van Harn (2004, p. 394)) whose characteristic function is $(1 + t^2)^{-\lambda} = (1 + it)^{-\lambda}(1 - it)^{-\lambda}$. Thus

$$V\sqrt{2\gamma(\lambda)} \stackrel{d}{=} \mathcal{L}(\lambda) := \gamma_1(\lambda) - \gamma_2(\lambda),$$

where $\gamma_1(\lambda) \perp \gamma_2(\lambda)$. The law of $\mathcal{L}(\lambda)$ has been called a generalised Laplace law because setting $\lambda = 1$ yields the Laplace law; see Kotz et al. (2001, p. 180). Now let N have the shifted Lerch law (9) and $N \perp \mathcal{L}(\lambda)$. It follows that if $X \sim GP(\lambda; a, c)$ with the parameter restrictions in Theorem 1, then

$$V\sqrt{2X} \stackrel{d}{=} \frac{\mathcal{L}(\lambda)}{N}$$

has the characteristic function $\Phi(c, \lambda, a + t^2)/\Phi(c, \lambda, a)$.

The following result shows that if λ is large then $L(X)$ is approximately normal. For this purpose we write $X(\lambda) = X$ to emphasise the dependence of its law on λ . A simple proof follows from Theorem 1 in the case that $c \geq 0$, but the proof we give in §10 for the general case follows from a probabilistic interpretation of the representation (5). Note that the outcome is independent of b and c .

Theorem 2 *If $X(\lambda) \sim GP(\lambda; a, b, c)$ then, as $\lambda \rightarrow \infty$,*

$$\frac{aX(\lambda) - \lambda}{\sqrt{\lambda}} \xrightarrow{d} N(0, 1), \tag{11}$$

the standard normal law.

The following result is an obvious consequence of (3).

Theorem 3 *If $X(\lambda) \sim GP(\lambda; a, b, c)$ then $X \xrightarrow{d} \gamma(\lambda)$ as $b \rightarrow \infty$.*

3 Some properties of the gamma law

The pdf of the general gamma random variable $\gamma(\lambda)/a$ is

$$g(x; \lambda, a) := g(x; \lambda, a, 0) = \frac{a^\lambda x^{\lambda-1} e^{-ax}}{\Gamma(\lambda)}, \quad (x > 0).$$

3.1 Unimodality

The gamma law is unimodal with mode

$$M_O = \begin{cases} 0 & \text{if } 0 < \lambda \leq 1, \\ \frac{\lambda-1}{a} & \text{if } \lambda > 1. \end{cases}$$

In addition, it is strongly unimodal if and only if $\lambda \geq 1$; see Dharmadhikari and Joag-Dev (1998, p. 23).

3.2 Infinite divisibility

The gamma law is infinitely divisible, in fact self-decomposable for all values of the shape parameter, $\lambda > 0$. See Steutel and van Harn (2004, p. 225).

3.3 Diffusion limiting law

Let $(Z_t : t \geq 0)$ be a diffusion process with $Z_0 \geq 0$, drift function $m(x)$, and diffusion function $\nu(x)$. Assume that these functions are such that the weak limit Z_∞ of (Z_t) exists with pdf $\zeta(x)$. See Horsthemke and Lefever (1984, p. 110) for these matters. In this case the integrated stationary forward Kolmogorov equation is

$$\frac{\zeta'(x)}{\zeta(x)} = 2 \frac{m(x)}{\nu(x)} - \frac{\nu'(x)}{\nu(x)}.$$

If $\nu > 0$ and $\nu(x) = 2x^\nu$, then the drift function is determined by $\zeta(x)$ as

$$m(x) = x^{\nu-1} \left(\nu + x \frac{\zeta'(x)}{\zeta(x)} \right) = x^{\nu-1} \left(\nu + x \frac{d}{dx} \log \zeta(x) \right). \quad (12)$$

In particular, if $\zeta(x)$ is the gamma pdf $g(x; \lambda, a)$, then

$$m(x) = x^{\nu-1}(\lambda + \nu - 1 - ax).$$

Feller (1951) solves the associated forward Kolmogorov equation in the case $\nu = 1$, thus justifying calling (Z_t) the *Feller diffusion*. This process arises in a demographic context as a continuous-time and continuous sample path analogue of the Galton-Watson branching process with immigration. In fact, under suitable restrictions, it emerges as a diffusion limit of such processes as the initial population size grows unboundedly and the individual reproduction mean tends to unity. See Li (2011, Chapter 3).

Incorporating mean reversion is accepted as a fundamental requirement for models of the short/spot rate of interest rates (Brigo and Mercurio (2001), p. 46). The Feller diffusion arises in quantitative finance theory, where it is named the Cox-Ingersoll-Ross (CIR) model, as a popular way of incorporating linear mean reversion. See Cox et al. (1985), and Lamberton and Lapeyre (1996, p. 131) for a monograph treatment. Its time varying one-dimensional distributions are derived in this reference. The fact that the gamma law is the limiting/stationary distribution seems not to be clearly documented, but see Wong (1984), and it is an obvious consequence of formulae on page 131 in Lamberton and Lapeyre (1996). The case $\nu = 2$ is a diffusion limit of the logistic model of population growth, for which see Horsthemke and Lefever (1984, §6.4) and references in Dennis and Patil (1984).

In general, and with a population growth interpretation, $r(x) = \lambda + \nu - 1 - ax$ is the per-capita rate of population growth which, if $\lambda + \nu > 1$, is positive if $x < (\lambda + \nu - 1)$ and it is negative if $x > (\lambda + \nu - 1)^+/a$. Equilibria are defined as positive solutions of $m(x) = 0$, and if $\lambda + \nu > 1$, there is one equilibrium, $x_e = (\lambda + \nu - 1)/a$.

Table 1 Classification of the boundary state $\{0\}$ for the diffusion model with $m(x) = x^{\nu-1}(\lambda + \nu - 1 - ax)$ and $v(x) = 2x^\nu$

Configuration	Attainability	Classification
$\lambda + \nu < 2$	$\{0\} \Leftrightarrow (0, \infty)$	Regular-reflecting
$\nu < 2 \leq \lambda + \nu$	$\{0\} \Rightarrow (0, \infty)$	Entrance
$\nu \geq 2$	$\{0\} \not\Rightarrow (0, \infty)$	Natural

Table 1 shows the dependence on the parameters (λ, ν) for the classification of the boundary state $\{0\}$. We use notation such as $\{0\} \Rightarrow (0, \infty)$ to mean that $(0, \infty)$ is attainable from $\{0\}$, but not the reverse. Note that this classification is independent of a .

3.4 Characterisation

The fundamental characterisation property of the gamma law (Lukacs (1965)) asserts that if X and Y are independent positive random variables such that $B = X/(X + Y)$ and $S = X + Y$ are independent, then X and Y have gamma laws with possibly different shape parameters, λ and μ , say. In this case $B \sim \text{beta}(\lambda, \mu)$ and $S \sim \text{gamma}(\lambda + \mu)$.

Many other characterisations are closely related to the Lukacs version. For example, we observe that length-biasing a gamma law yields another gamma law. More exactly (and recalling the length-bias operator \mathcal{L}_r), if $r > 0$, then

$$\mathcal{L}_r \gamma(\lambda) \stackrel{d}{=} \gamma(\lambda + r). \tag{13}$$

Hence one direction of the Lukacs characterisation can be expressed as a fixed-point property: If $B \sim \text{beta}(\lambda, r)$ is independent of $\gamma(\lambda)$, then $\gamma(\lambda) \stackrel{d}{=} B\mathcal{L}_r \gamma(\lambda)$. This is a characteristic property in the following sense (Pakes (1997, Theorem 4.1)).

Lemma 1 *Let $\lambda, r > 0$. Any two of the following implies the third:*

- (a) $Y \stackrel{d}{=} V\mathcal{L}_r Y$ and $E(Y^r) = \lambda$;
- (b) $L(V) = \text{beta}(\lambda, r)$; and
- (c) $L(Y) = \text{gamma}(\lambda)$.

4 Planck laws as diffusion stationary laws

Let $\zeta(x)$ be the pdf of the gamma diffusion limit law in §3. Following Dennis and Patil (1984), let $w(x)$ be a positive weight function such that $\int_0^\infty w(x)\zeta(x)dx < \infty$. It follows from (12) that the pdf $\propto w(x)\zeta(x)$ is that of the limiting/stationary law for the diffusion with $v(x) = 2x^\nu$ and infinitesimal mean function

$$m_w(x) = m(x) + x^{\nu-1} \frac{w'(x)}{w(x)}.$$

The $GP(\lambda; a, b, c)$ pdf (3) is the weighted gamma pdf with $w(x) = (1 - ce^{-bx})^{-1}$, whence

$$m_w(x) = \left(\lambda + \nu - 1 - ax - \frac{bcx}{e^{bx} - c} \right) x^{\nu-1}.$$

The perturbation term $p(x) = bcx/(e^{bx} - c)$ is positive and bounded if $0 \leq c \leq 1$ and it is negative and bounded if $c < 0$. If $c = 1$ then $p(x)$ decreases from unity at the origin to zero as $x \rightarrow \infty$. If $c < 1$ then $p(x)$ has a single critical value where $bx = y(c)$ solves $(1 - y)e^y = c$, that is, $y(c) = 1 + W(-c/e)$ where W is the principal Lambert W -function; see Olver et al. (2010, #4.13). In particular $y(c)$ increases as c decreases, $y(1) = 0, y(0) = 1$

and $1 < y(c) \uparrow \infty$ as $c \downarrow -\infty$. The extreme value of the perturbation function is $1 - y(c)$. In particular, if $\nu = 2$ then the logistic per-capita growth rate $r(x)$ is perturbed by a localised ‘bump’ which alters the limiting/stationary law in a quite significant manner.

In general, the perturbation term lessens the equilibrium level, $x_{e,w} < x_e$ (defined in §3.4), though much less so if x_e is large. The rate of mean reversion is less for $x < x_{e,w}$, but if $c = 1$ then the rate near the origin is close to $\lambda + \nu - 1$. The classification of $\{0\}$ in Table 1 is unchanged if $c < 1$, but if $c = 1$ then the table entries are valid if λ is replaced by $\lambda - 1$ (which is positive in this case).

One aim of Dennis and Patil (1984) is exhibiting perturbations of the Feller/CIR diffusion whose limit laws are multi-modal, at least for certain parameter combinations. Their motivation is constructing stochastic versions of one-dimensional dynamical systems which have several equilibria. We examine modal behaviour of the GP laws in the next section.

5 Modal properties of GP laws

Nadarajah and Kotz (2006) list parameter conditions under which a Planck pdf is decreasing, unimodal (i.e., a unique positive mode), or ‘maybe’ multi-modal. The modal behaviour of GP laws also is of interest in statistical physics, e.g., Valluri et al. (2009) and references. We give here more precise statements, beginning with the more easily answered question of strong unimodality.

5.1 Strong unimodality

Recall that a pdf $g(x)$ is strongly unimodal if and only if g is log-concave (Dharmadhikari and Joag-Dev (, p. 17)). We work with the GP pdf $g(x) = g(x; \lambda, a, c)$ and since this is twice continuously differentiable, then this condition is equivalent to

$$C(x) := \frac{d^2}{dx^2}(\log g(x)) = \frac{-\ell}{x^2} + \frac{ce^x}{(e^x - c)^2} \leq 0, \quad (x > 0), \tag{14}$$

where, for convenience, we write $\ell = \lambda - 1$. Thus $\ell > 0$ if $c = 1$ and $\ell > -1$ if $c < 1$. The following result is specified in terms of the shape parameter λ . We write $g \in SU$ to denote that g is strongly unimodal.

Theorem 4 *Let $g(x) = g(x; \lambda, a, c)$ with $\lambda > 0$ and $-\infty < c \leq 1$, but $c \neq 0$.*

- (i) $g \notin SU$ if $0 < c < 1$ and $0 < \lambda < 1$.
- (ii) If $c < 0$ then $g \in SU$ if and only if $\lambda \geq 1$.
- (iii) If $c = 1$ and $\lambda > 1$, then $g \in SU$ if and only if $\lambda \geq 2$.
- (iv) Let $0 < c < 1$ and $\varepsilon = (1 - c)/(1 + c)$. Then there exists a unique number $\Upsilon \in (\varepsilon, 1)$ such that

$$e^{2\Upsilon} = c \frac{1 + \Upsilon}{1 - \Upsilon} \tag{15}$$

and $g \in SU$ if and only if $\lambda \geq 2 - \Upsilon^2$.

5.2 Unimodality

Strong unimodality implies unimodality, but the two notions differ and the algebraic work for establishing the latter is not as clean cut as for the former. In this subsection we work with the full family $GP(\lambda; a, b, c)$, recalling that $\alpha = a/b$. Denote the pdf by

$g(x) := g(x; \lambda, a, b, c)$. The following preliminary result is a simple deduction from (3). Proofs of the following results are in §10.

Lemma 2 *Values of g at the origin are*

$$g(0+) \begin{cases} = 0 & \text{if } c = 1 \text{ and } \lambda > 2, \text{ or if } c < 1 \text{ and } \lambda > 1; \\ \in (0, \infty) & \text{if } c = 1 \text{ and } \lambda = 2, \text{ or if } c < 1 \text{ and } \lambda = 1; \\ = \infty & \text{if } c = 1 \text{ and } 1 < \lambda < 2, \text{ or if } c < 1 \text{ and } 0 < \lambda < 1. \end{cases}$$

A calculation yields

$$\Lambda_g(x) := \frac{g'(x)}{g(x)} = b \left(\frac{\ell}{bx} - \alpha - \frac{c}{e^{bx} - c} \right). \tag{16}$$

Setting $y = bx$, we see that the sign of $g'(x)$ is determined by the sign in $(0, \infty)$ of

$$S(y) = (\ell - \alpha y) (e^y - c) - cy. \tag{17}$$

We begin with the case of most physical significance, $\alpha = c = 1$.

Theorem 5 *If $\alpha = c = 1$, then the pdf $g(x) := g(x; \lambda, a, b, c)$ is:*

- (i) *Decreasing if $1 < \lambda \leq 2$; and*
- (ii) *Unimodal if $\lambda > 2$. In this case the mode is*

$$M_O = a^{-1}[\lambda - 1 + W(-(\lambda - 1)e^{-(\lambda-1)})] > 0,$$

where $W(\cdot)$ is the principal Lambert function.

- (iii) *If $\lambda \gg 1$, then*

$$M_O = \frac{\lambda - 1}{a} \left(1 - e^{-(\lambda-1)} \right) (1 + o(1)).$$

Remark 2 *Spectroscopists define the normalised emissive power of a radiation distribution as the spectral density function divided by its modal value. Under the conditions of Theorem 5, this has the explicit form*

$$\frac{g(x)}{g(M_O)} = \frac{(x/M_O)^{\lambda-1}}{[e^{ax} - 1] [((\lambda - 1)/aM_O) - 1]}.$$

See Stewart (2012) (and its references) for the Planck distribution.

The next result relaxes the condition on α .

Theorem 6 *If $c = 1$, then the pdf $g(x) := g(x; \lambda, a, b, c)$ is:*

- (i) *Unimodal and $g(0) = 0$ if $\lambda > 2$;*
- (ii) *If $1 < \lambda < 2$, then $g(0) = \infty$ and there is a critical value α_c of $\alpha = a/b$ such that $g(x)$ decreases if $\alpha \geq \alpha_c$. If $\alpha < \alpha_c$, then $g(x)$ has an anti-mode (i.e., a local minimum) M_A and a mode M_O satisfying $0 < M_A < M_O < \lambda - 1$.*
- (iii) *If $\lambda = 2$, then $0 < g(0) < \infty$ and the modal behaviour of (ii) holds with $\alpha_c = \frac{1}{2}$.*

Example 1 *Nadarajah and Kotz (2006) suggest that g ‘may’ be multi-modal if $a = c = 1$, $b = 10$ and $\lambda = 1.75$. Referring to the Proof of Theorem 6, the unique positive solution \bar{z} of $z^{-1} \sinh z = \ell^{-\frac{1}{2}}$ is $\bar{z} = 1.88461$ whence $\alpha_c = 0.21887$. But $\alpha = 0.1$, so g is bimodal. The two solutions of (49) are 0.7372 and 7.74569, whence $M_A = 0.07372$ and $M_O = 0.7457$.*

Next, we let $c < 1$ and distinguish the two sub-cases where $0 < c < 1$ and $c < 0$. In both cases the parameter $\ell \in (-1, \infty)$. In the case $0 < c < 1$, if $0 < \lambda \leq 1$, then $S(y) < 0$ in $(0, \infty)$, and if $\lambda > 1$, then $S(y) < 0$ in $(\ell/\alpha, \infty)$. These observations imply the first two assertions of the following result.

Theorem 7 Let $0 < c < 1$.

- (i) If $0 < \lambda \leq 1$, then $g(x)$ is decreasing.
- (ii) If $\lambda > 1$, then critical values of $g(x)$, if they exist, are confined to $(0, (\lambda - 1)/a)$.
- (iii) If $\lambda > 1$ and $\alpha \geq \frac{1}{2}(\lambda - 1)$, then $g(0) = 0$ and $g(x)$ is unimodal.
- (iv) If $\lambda > 1$ and $\alpha < \frac{1}{2}(\lambda - 1)$, then $g(x)$ is either unimodal or bi-modal.

Example 2 To see that bi-modality really is possible, we use parameter values based on those cited with no definite conclusion by Nadarajah and Kotz (2006): $\lambda = 3/2$, $a = 1$, $b = 10$ and $c = 0.9$. Table 2 displays the three critical values x_i of the pdf, i.e. zeros of $S(bx)$, and values of its un-normalised version i.e., referring to (3), $\widehat{g}(x) = x^{\lambda-1}e^{-ax}(1 - ce^{-bx})^{-1}$.

Theorem 8 Assume that $c < 0$. Then the pdf $g(x) = g(x; \lambda, a, b, c)$ is:

- (i) Unimodal if $\lambda > 1$ and $M_O > 0$;
- (ii) Unimodal if $\lambda = 1$ and then

$$M_O = \begin{cases} 0 & \text{if } \alpha \geq (-c)/(1 - c), \\ \log(-c(\alpha^{-1} - 1)) & \text{if } \alpha < (-c)/(1 - c) (< 1); \end{cases}$$

- (iii) If $0 < \lambda < 1$, then $g(0+) = \infty$ and there is a critical value $\lambda_c = 1 - W(-ce^{-1})$ such that
 - (iiia) If $\lambda_c \leq \lambda < 1$, then $g(x)$ is decreasing for all $\alpha > 0$ and the condition $c \geq -(1 - \lambda)e^{2-\lambda}$ is necessary for this outcome, and
 - (iiib) If $0 < \lambda < \lambda_c$, then there is a critical value α_c such that $g(x)$ is decreasing if $\alpha > \alpha_c$, and if $0 < \alpha < \alpha_c$, then there is an anti-mode M_A and a mode M_O such that $0 < M_A < M_O$.

Example 3 Again following Nadarajah and Kotz (2006), let $\lambda = 0.99$, $a = 1$, $b = 10$ and $c = -1/3$. Using the iteration $W_{n+1} = ye^{-W_n}$ to compute $W(y)$, it quickly follows that $W(-ce^{-1}) = 0.10987 = R(\bar{y})$. Hence $\lambda_c = 0.09013$ and the solutions of (50) are 0.07325 and 0.99123. It follows that $M_A = 0.007325$ and $M_O = 0.099122$.

6 Infinite divisibility

We begin by proving a mixture representation for the case $c > 0$ which extends Theorem 1 and which can be regarded as an example of the general principle which can

Table 2 Modes are $M_{O1} = x_1$ and $M_{O2} = x_3$ and the anti-mode $M_A = x_2$

x_i	$\widehat{g}(x_i)$
$x_1 = 0.012813$	$\widehat{g}(x_1) = 0.536670$
$x_2 = 0.174431$	$\widehat{g}(x_2) = 0.416273$
$x_3 = 0.456995$	$\widehat{g}(x_3) = 0.432069$

be inferred from Stuart (1962) and which is explicit in Steutel and van Harn (2004, Proposition 4.2, p. 344). This principle asserts that if $0 < \lambda < \nu$, then a scale mixture of the $\text{gamma}(\lambda)$ law can be expressed as a scale mixture of the $\text{gamma}(\nu)$ law. We write the fundamental characterisation as $\gamma(\lambda) \stackrel{d}{=} B\gamma(\nu)$ where $B \sim \text{beta}(\lambda, \nu - \lambda)$ (with the convention that $P(B = 1) = 1$ if $\nu = \lambda$) and $B \perp N$. The following mixture representation follows from Theorem 1.

Theorem 9 *Let $a, \lambda > 0, 0 \leq c \leq 1$ and $X \sim GP(\lambda; a, c)$. Then for each $\nu \geq \lambda, X \stackrel{d}{=} \gamma(\nu)/V$ where*

$$V = N/B, \quad B \sim \text{beta}(\lambda, \nu - \lambda), \quad B \perp N, \quad \& \quad \gamma(\nu) \perp V, \tag{18}$$

and N has the shifted Lerch law as in Theorem 1. The pdf of V is

$$f_V(\nu) = \begin{cases} [B(\lambda, \nu - \lambda)\Phi(c, \lambda, a)]^{-1} \nu^{-\nu} \sum_{n \leq \nu - a} c^n (\nu - a - n)^{\nu - \lambda - 1} & \text{if } \nu > a, \\ 0 & \text{if } \nu < a. \end{cases} \tag{19}$$

The evaluation (19) results by differentiating the identity $P(V > \nu) = \sum_n P(B > (a + n)/\nu)p(n)$ to get

$$f_V(\nu) = \nu^{-2} \sum (a + n)f_B((a + n)/\nu)p(n),$$

substituting the algebraic form of the beta pdf and observing that $f_B((a + n)/\nu) = 0$ if $\nu < a$ or if $\nu > a$ and $n > \nu - a$.

The nature of $f_V(\nu)$ changes quite drastically as ν decreases to λ . If $\nu - \lambda > 1$, this pdf is continuous; if $\nu - \lambda = 1$ then it is decreasing with jump discontinuities everywhere on the support of N ; if $0 < \nu - \lambda < 1$ then this pdf has singularities on the support of N . Clearly $V \xrightarrow{L} N$ as $\nu \downarrow \lambda$.

If $0 < c < 1$ and $X \sim GP(\lambda; a, c)$, then the representation assertion of Theorem 9 can be expressed as $X \stackrel{d}{=} BW$ where $W = \gamma(\nu)/N$ whose pdf is

$$f_W(w) = (\Gamma(\nu))^{-1} w^{\nu-1} e^{-aw} \cdot \frac{\Phi(ce^{-w}, -(v - \lambda), a)}{\Phi(c, \lambda, a)}.$$

This defines a family of laws which includes the $GP(\lambda; a, c)$ family as a special case; just observe that $\Phi(c, 0, a) = 1/(1 - c)$.

In what follows we use the abbreviation *infdiv* to mean infinitely divisible. It follows from Theorem 9 that the $GP(\lambda; a, b, c)$ law is a mixture of any $\text{gamma}(\nu)$ law with $\nu \geq \lambda$. Mixtures of $\text{gamma}(2)$ laws are *infdiv* (Steutel and van Harn (2004, p. 346)), implying Assertion (i) in the following theorem. Similarly, any parameter configuration allowing the choice $\nu = 1$ in Theorem 9 gives $L(X)$ in the form of a mixture of exponential distributions and hence, as is obvious from (3), the pdf $g(x)$ is completely monotone and hence convex. In fact, inspection of (14) shows that $g(x)$ is log-convex if $\lambda \leq 1$ and $0 < c \leq 1$. Log-convexity of a pdf is a sufficient condition that it be *infdiv* (e.g. Sato (2013, Theorem 51.4)) (and hence not strongly unimodal). Assertion (ii), stated for the case $b = 1$, extends a little the ambit of Assertion (i). Its proof is in §10.

Theorem 10 (i) *If $0 \leq c \leq 1$ and $\lambda \leq 2$, then the $GP(\lambda; a, b, c)$ law is *infdiv*.*

(ii) *Let $c < 0$. The $GP(\lambda; a, b, c)$ law is *infdiv* if*

$$-e^2 = -7.389056 < c < 0 \quad \text{and} \quad \lambda \leq 1 - W(-ce^{-1}). \tag{20}$$

Remark 3 Let $v > 0$ and observe that ve^{v+1} increases with v . Observing that $W(ve^v) = v$, the left-hand constraint in (20) is equivalent to $0 < v < 1$ in which case $L(X)$ is infdiv if $0 < \lambda < 1 - v$.

Remark 4 The conditions (20) are broader than those under which the pdf $g(x; \lambda, a, b, c)$ is log-convex, a sufficient condition for infinite divisibility. Let $c' = -(3 - 2\sqrt{2})e^{2\sqrt{2}} = -2.902812$. Still assuming that $c < 0$, then g is log-convex if and only if $c' < c < 0$ and $0 < \lambda \leq \lambda_c$, where $\lambda_c \in (0, 1)$ is a certain critical value. Details for this assertion comprise part of the Proof of Proposition 1 (stated below). For example, if $c = -1$, then $\lambda_c = 0.56077$.

It follows from Theorem 10 that the Bose-Einstein laws ($\lambda = 1$) are infdiv, but the status of the Planck law ($\lambda = 4$) is unknown. In passing we mention Varrò (2007) whose Section 5 is titled ‘The infinite divisibility of the Planck variable’. His discussion distills to the assertion of the well-known fact that the geometric law is infdiv - he ascribes the title ‘Planck-Bose’ to this law.

Suppose that $0 < c \leq 1, \lambda \leq 2$ and that $b = 1$, so $X \sim GP(\lambda; a, c)$ is infdiv. Note that the corresponding Lévy process is an infinite-activity (Type 1) subordinator, i.e., its sample paths are non-decreasing. It seems to be difficult to determine the nature of the Lévy measure of $L(X)$. We can obtain some insight as follows. It is clear from its definition that the distribution family $\{GP(\lambda; a, c) : a > 0\}$ is the natural exponential family generated by the measure whose density is

$$\beta(x) = \frac{x^{\lambda-1}}{\Gamma(\lambda)(1 - ce^{-x})}$$

with Laplace transform

$$\widehat{\beta}(\theta) = \Phi(c, \lambda, \theta) = e^{-C(\theta)},$$

where $C(\theta) = -\log \Phi(c, \lambda, \theta)$ is a Bernstein function, i.e., $C'(\theta)$ is completely monotone (Schilling et al. (2012)). Thus there is a measure which, on the basis of the following discussion, we conjecture is absolutely continuous with density denoted by $\ell(x)$, such that

$$C'(\theta) = \int_0^\infty e^{-\theta x} x \ell(x) dx. \tag{21}$$

It follows that the Laplace exponent of $L(X)$ is $C(a + \theta) - C(a)$ and its Lévy measure is $e^{-ax} \ell(x) dx$.

Next, we have

$$C'(\theta) = -\frac{(\partial/\partial\theta)\Phi(c, \lambda, \theta)}{\Phi(c, \lambda, \theta)} = \frac{\lambda\Phi(c, \lambda + 1, \theta)}{\Phi(c, \lambda, \theta)}. \tag{22}$$

Proceed formally by writing $C'(\theta) = \sum_{n=0}^\infty c^n k_n(\theta)$ and equate the coefficients of c^n in the identity obtained from (22) by multiplying through by $\Phi(c, \lambda, \theta)$ and using the representation (4). This yields

$$\sum_{j=0}^n \frac{k_j(\theta)}{(n - j + \theta)^\lambda} = \frac{\lambda}{(n + \theta)^{\lambda+1}},$$

allowing the step-by-step determination of the $k_n(\theta)$. Thus $k_0(\theta) = \lambda/\theta$, the Laplace transform of the measure λdx . This yields the Lévy measure $(\lambda/x)e^{-ax} dx$ for the case $c = 0$, i.e., for the gamma laws. Continuing, we find that

$$k_1(\theta) = \frac{-\lambda}{\theta^{1-\lambda}(1+\theta)^{1+\lambda}}$$

and

$$k_2(\theta) = \frac{\lambda}{\theta^{1-2\lambda}(1+\theta)^{1+2\lambda}} - \frac{2\lambda}{\theta^{1-\lambda}(2+\theta)^{1+\lambda}}.$$

Denoting the Kummer confluent hypergeometric function (Olver et al. (2010, p. 322)) by $M(\psi, \rho, x)$, it follows easily from an integral representation that the Laplace transform of $xM(1+a, 2, -vx)$ is $\theta^{1-a}(v+\theta)^{-1-a}$. Thus $k_i(\theta)$ ($i = 1, 2$) can formally be inverted to give the expansion

$$\ell(x) = \lambda(x^{-1} - cM(1+\lambda, 2, -x) + c^2\{M(1+2\lambda, 2, -x) - M(1+\lambda, 2, -2x)\}) + O(c^3).$$

However, the forms of $k_n(\theta)$ for $n \geq 3$ become more complicated and the corresponding inverse transforms involve convolutions of algebraic and Kummer functions with more and more factors.

We remind the reader that X has a self-decomposable (SD) law if for all $\rho \in (0, 1]$ it has the auto-regression representation $X \stackrel{d}{=} \rho X + Y_\rho$, where $Y_\rho \perp X$. Self-decomposable laws are infdiv and they arise precisely as the possible weak limits of normed sums of independent random variables. We have seen that Planck laws can be bimodal under certain interacting parameter combinations, and since self-decomposable laws are unimodal, we may expect similar interacting parameter combinations to appear in the following partial result on self-decomposability of Planck laws. Its proof is in §10. Recall the constant c' from Remark 4.

Proposition 1 *If $0 < \lambda \leq 1$ then the GP($\lambda; a, b, c$) law is self-decomposable if:*

- (i) $c = 1$ and $\alpha \geq \frac{1}{2}$; or
- (ii) $0 < c < 1$ and $\alpha \geq \alpha_c := \frac{1}{2}(1 - \frac{1}{2}y_c)$, where $y_c \in (0, 2)$ is the unique positive solution of

$$e^y = c \frac{2+y}{2-y}; \tag{23}$$

or

- (iii) $c' < c < 0$, $\alpha \geq \alpha_c := -c/(1-c)$, and $0 < \lambda \leq \lambda_c := 2 - y_c^2/4$, where $y_c \in (2, \infty)$ is the unique positive solution of (23).

Remark 5 *Observe first that $y_c \rightarrow 2$ as $|c| \rightarrow 0$ and hence the upper bound for λ in part (iii) is positive if c is close to zero. More precisely, suppose that $0 < |c| < 1$, write $y_c = 2 - \delta$ and observe that $c\delta > 0$. It follows from (23) that, as $c \rightarrow 0$,*

$$y_c = 2 - 4ce^{-2} + o(c^2).$$

So, for $c \rightarrow 0+$, the bound for Assertion (ii) is $\alpha \geq ce^{-2}(1 + o(c))$, and for $c \rightarrow 0-$, the bound for Assertion (iii) is $\lambda \leq 1 + 4ce^{-2}(1 + o(1))$.

Remark 6 *The bound for y_c can be improved. If $0 < c < 1$, it follows from (23) and $y_c < 2$ that $(y_c - 2)e^{y_c} = -(2 + y_c)c > -4c$. Hence*

$$y_c > 2 + W(-4ce^{-2}).$$

This lower bound subsists if $c < 0$ and $\lambda_c = 0$ if $c = c'$.

Table 3 Representative critical values for self-decomposability in Proposition 1

c	1	0.5	-0.5	-1	-1.5	-2	-2.5	-2.9
γ_c	0	1.6493	2.2278	2.3994	2.5379	2.6547	2.7557	2.8279
α_c	1/2	0.0877	0.3333	0.5	0.6	0.6667	0.7173	0.7436
λ_c	1	1	0.7592	0.5608	0.3897	0.2382	0.1015	0.000685

Example 4 Calculation yields the representative values of σ_c and λ_c displayed in Table 3.

In the self-decomposable case of Proposition 1, the random variable X has a stochastic integral representation which we express as

$$X = \int_0^\infty e^{-t} dL_t,$$

where $(L_t : t \geq 0)$ is a Lévy process, which, in this case has non-decreasing sample paths (i.e., is a subordinator), and called the background driving Lévy process (conventionally abbreviated to BDLP); see Wolfe (1982). The Laplace exponent of the BDLP is $C_b(\theta) = \theta C'(\theta)$. It follows from (22) that

$$C_b(\theta) = \frac{\lambda\theta\Phi(c, \lambda + 1, a + \theta)}{\Phi(c, \lambda, a + \theta)}, \tag{24}$$

and hence from (43) below that

$$C_b(\theta) = \frac{\lambda\theta}{1 + \theta} (1 + o(1)) \quad (\theta \rightarrow \infty).$$

In particular, $C_b(\infty-) = \lambda$, implying that the BDLP is a compound Poisson process. We infer too from Proposition 1 that the quotient (24) is a Bernstein function at least for the range of c delineated in the assertion.

We have the following consequence, implying that X is the quotient of independent infdiv random variables.

Theorem 11 *Let $0 \leq c < 1$ and $\lambda > 0$ or $c = 1$ and $\lambda > 1$. Then $L(N)$ is infdiv.*

The following result is a statement about what is sometimes called multiplicative infinite divisibility. Its proof is in §10.

Lemma 3 *If $c = 1, \lambda > 1$ and $b = a$ or $b = 2a$, then $\log X$ is infdiv.*

What if $c < 1$? The answer is not known except when $c = 0$, the log-gamma law, $\Lambda := -\log \gamma(\lambda)$. It follows from Euler’s product representation of the gamma function that

$$\sum_{j=1}^n \frac{\gamma_j(1)}{j + \lambda - 1} - \log n \xrightarrow{d} \Lambda.$$

This exhibits Λ as the limit of a centred sum of independent gamma distributed random variables. The set of all such limit laws comprise the class of extended generalised gamma convolutions and they are SD laws. See Bondesson (1992, p. 112) and Steutel and van Harn (2004, p. 322).

We remark in passing that a direct corollary of the derivation of this result is that

$$\sum_{j=1}^n \gamma_j(\lambda)/j - \lambda \log n \xrightarrow{d} Z_\lambda$$

has an extended generalised gamma convolution law with

$$E(\exp(-tZ_\lambda)) = (\Gamma(1 + t))^\lambda. \tag{25}$$

This shows in particular that the sequence $\{n^\lambda : n = 0, 1, \dots\}$ is a Stieltjes moment sequence. Berg (2005, §2) has discussed this moment sequence in detail, showing in particular that it is Stieltjes-determinate if and only if $\lambda \leq 2$. (Note that Berg denotes our parameter λ by c .)

Observing that $\log(n + 1) - \log n \rightarrow 0$ ($n \rightarrow \infty$), we have the random series representation

$$Z_\lambda \stackrel{d}{=} \sum_{j=1}^\infty (j^{-1} \gamma_j(\lambda) - \lambda \log(1 + 1/j)).$$

Hence (25) is equivalent to $E(B_\lambda^t) = (\Gamma(1 + t))^\lambda$, where

$$B_\lambda = \prod_{j=1}^\infty (1 + 1/j)^\lambda \exp(j^{-1} \gamma_j(\lambda)) \tag{26}$$

has the pdf $e_\lambda(x)$ identified in Lemma 2.1 of Berg (2005). Clearly $B_\lambda \times B_\nu \stackrel{d}{=} B_{\lambda+\nu}$, thus providing another proof that the family of pdf's $(e_\lambda : \lambda > 0)$ comprises a product convolution semigroup of pdf's.

We note that Berg's apparently different analytical proofs of his results derive in part from a product representation of the gamma function.

We now recall that the weak limits of sums of independent gamma random variables (i.e., no centring) are called generalised gamma convolutions (GGC's) and they inherit the self-decomposability of gamma laws. Thus, the GGC class is closed under independent addition of random variables and, remarkably, it also is closed under independent multiplication (Bondesson (2015)).

The discrete-law version of GGC's are the generalised negative binomial convolution's (GNBC's), defined as the weak limits of sequences of sums of independent negative binomial random variables. The GNBC's coincide with the mixtures of Poisson laws where the mixing law is a GGC. Hence GNBC laws are infdiv. The next result records the known fact that a Lerch law is a GNBC; see Bondesson (1992, p. 135) or Steutel and van Harn (2004, p. 420). The proofs in these references are distributed over widely separate portions of their expositions, so it is worthwhile to record a consolidated proof. This proof is in §10 together with necessary concepts and facts.

Theorem 12 *If $0 < c \leq 1$, then the Lerch law $L(N_L)$ is a GNBC if: (i) $\alpha > 0$ and $\lambda \geq 1$, or (ii) $\alpha > 1$ and $\lambda > 0$.*

Remark 7 *A referee of an earlier version of this paper observed that the GNBC property may not hold if $0 < \alpha, \lambda < 1$, and indeed it does not if $\alpha = 0.5$ and $\lambda = 0.25$.*

It follows from Theorems 1 and 12 that X is equal in law to $\gamma(\lambda)$ divided by an independent GNBC random variable. However this is not sufficient to conclude that $L(X)$ is

infdiv. Recent success in proving that many common laws are infdiv follows from showing that they are GGC's. An important proper subset of GGC's comprise absolutely continuous laws whose pdf's are hyperbolically completely monotone (HCM). A function $H(x)$ is called HCM if, for each $u > 0$, the product $H(uv)H(u/v)$ is a completely monotone function of $w = v + v^{-1}$ where $v > 0$. See Bondesson (1992) or Steutel and van Harn (2004, §5 in Chap. VI). Any gamma pdf is HCM and many other pdf's have this property. However, we have the following counter-result.

Theorem 13 *The pdf $g(x; \lambda, a, b, c)$ is HCM if and only if $c = 0$*

Proof An HCM function extends as a holomorphic function to the complex plane cut along the negative reals (Bondesson (1992, p. 68)). But clearly, if $c \neq 0$ and $\rho := -\log |c|$, the denominator of the pdf (3) has simple poles at

$$z_n = \begin{cases} -b^{-1}(\rho + 2ni\pi) & \text{if } c > 0, \\ -b^{-1}(\rho + (2n - 1)i\pi) & \text{if } c < 0, \end{cases} \quad n = 0, \pm 1, \dots$$

The assertion follows. □

This outcome leaves open the possibility that a GP law is a GGC, but the following argument suggests not. Suppose that $0 < |c| \leq 1$. It follows from (7) that the moment generating function of the GP law is proportional to $m(s) = \Phi(c, \lambda, \alpha - s)$ and which, by virtue of (4), is analytic in the complex plane except for singularities at $\alpha, \alpha + 1, \dots$. If the GP law is a GGC, then for $s > 0$ its Thorin measure $U(ds)$ is computed as $U((0, s)) = \pi^{-1}A(s)$ where $A(s) := \arg(m(s))$; see Bondesson (1992, p. 31). In particular $A(s)$ is non-decreasing in $(0, \infty)$. Clearly $A(s) \equiv 0$ in $(0, \alpha)$. Using the evaluation $-1 = e^{-i\pi}$, where $i = \sqrt{-1}$, we see for $\alpha < s < \alpha + 1$ that

$$m(s) = (s - \alpha)^{-\lambda} e^{i\pi\lambda} + \sigma(s)$$

where $\sigma(s) = \sum_{n \geq 1} c^n (\alpha + n - s)^{-\lambda}$ is real and positive-valued in this interval if $0 < c \leq 1$, and negative-valued if $-1 \leq c < 0$. Hence

$$\tan A(s) = \frac{\sin \pi \lambda}{\cos \pi \lambda + (s - \alpha)^\lambda \sigma(s)}.$$

This is identically zero if λ is a natural number, so assume this is not the case. Letting $s \downarrow \alpha$ shows that $U(ds)$ assigns mass λ to $s = \alpha$, exactly as for a gamma law, the case $c = 0$. However $|\sigma(s)| \uparrow \infty$ as $s \uparrow \alpha + 1$ and hence $\tan A(s) \rightarrow 0$. Consequently $A(s)$ cannot be non-decreasing in $(\alpha, \alpha + 1)$, and this contradiction implies that the $GP(\lambda; a, b, c)$ law is a GGC if and only if $c = 0$.

7 Characterisations by weighting and random scaling

In this section, factors in products of random variables are assumed to be independent. As mentioned in §1, many laws $L(Y)$ have been characterised in terms of a fixed point relation expressed as $Y \stackrel{d}{=} V\mathcal{L}_r Y$, e.g., Pakes (1997). We begin by seeking to extend Lemma 1 to the Planck laws. We approach this by observing that if $Z \geq 0$ is independent of Y and $E(Z^r) < \infty$, then $\mathcal{L}_r(YZ) \stackrel{d}{=} (\mathcal{L}_r Y)(\mathcal{L}_r Z)$, so if $U\mathcal{L}_r Z \stackrel{d}{=} Z$ then $UV\mathcal{L}_r(YZ) \stackrel{d}{=} YZ$. Hence characterising a product law such as $GP(\lambda; a, b, c)$ can be achieved by attending separately to its factors. The gamma factor in Theorem 1 is addressed by Lemma 1, so we need only to consider the shifted Lerch zeta law divisor N .

We will find it useful to highlight the dependence of $L(N)$ on λ by writing $N(\lambda)$ for N . It follows from (10) that the moment function of $\mathcal{L}_{-r}N(\lambda)$ is

$$\frac{E(N^{-r-t}(\lambda))}{E(N^{-r}(\lambda))} = \frac{\nu(r+t)}{\nu(r)}.$$

According to (10) the right-hand side equals $EN^{-t}(\lambda+r)$, and hence

$$\mathcal{L}_{-r}N(\lambda) \stackrel{d}{=} N(\lambda+r), \tag{27}$$

a property which is reciprocal to (13).

Let \mathcal{P} denote the set of prime numbers and let $\{N_p\}$ be a sequence of independent random variables having the ‘stretched’ geometric laws

$$P(N_p = p^n) = (1 - p^{-\lambda})p^{-n\lambda}, \quad (n \geq 0). \tag{28}$$

Clearly

$$\mu_p(t) := E(N_p^{-t}) = \frac{1 - p^{-\lambda}}{1 - p^{-\lambda-t}}. \tag{29}$$

Next, define independent random variables $U_{p,r}$ ($p \in \mathcal{P}$) such that

$$P(U_{p,r} = p^n) = \begin{cases} m(p,r) := (1 - p^{-\lambda}) / (1 - p^{-\lambda-r}) & \text{if } n = 0, \\ m(p,r) (1 - p^{-r}) p^{-n\lambda} & \text{if } n = 1, 2, \dots \end{cases} \tag{30}$$

Finally, we state a useful general property of the length bias operator which extends Lemma 2.1 (c) in Pakes (1997).

Lemma 4 *If Y is a positive random variable and $r > 0$, then $\mathcal{L}_r Y^s \stackrel{d}{=} (\mathcal{L}_{rs} Y)^s$ for any real s such that $E(Y^{rs}) < \infty$.*

Proof The definition of \mathcal{L}_r yields

$$E[(\mathcal{L}_r Y^s)^t] = \frac{E[Y^{s(t+r)})]}{E[Y^{rs}]} = \frac{E[Y^{rs+st}]}{E[Y^{rs}]} = E[(\mathcal{L}_{rs} Y)^{st}].$$

□

The following result is effectively a characterisation of $L(N_p)$ in terms of an inverse power-law weighting.

Lemma 5 *If $\lambda, r > 0$ then*

$$N_p \stackrel{d}{=} U_{p,r} \mathcal{L}_{-r} N_p.$$

Proof It follows from the evaluation (29) and using (30) to compute

$$E(U_{p,r}^{-t}) = m(p,r) \frac{1 - p^{-\lambda-rt}}{1 - p^{-\lambda-t}}$$

that

$$E\left[\left(\mathcal{L}_r N_p^{-1}\right)^t\right] = \frac{\mu_p(t+r)}{\mu_p(r)} = \frac{1 - p^{-\lambda-r}}{1 - p^{-\lambda}} \cdot \frac{1 - p^{-\lambda-t}}{1 - p^{-\lambda-r-t}} \mu_p(t) = \frac{\mu_p(t)}{E[U_{p,r}^{-t}]}.$$

This identity implies that

$$N_p^{-1} \stackrel{d}{=} U_{p,r}^{-1} \mathcal{L}_r N_p^{-1}. \tag{31}$$

Hence $N_p \stackrel{d}{=} U_{p,r}(\mathcal{L}_r N_p^{-1})^{-1} \stackrel{d}{=} U_{p,r} \mathcal{L}_{-r} N_p$, where the last step follows from Lemma 7.2 with $s = -1$. \square

Now assume that $\alpha = c = 1$ in which case the moment formula (10) becomes $v(t) = \zeta(\lambda + t)/\zeta(\lambda)$. The Euler product formula and (28) yield the product representation

$$N(\lambda) \stackrel{d}{=} \prod_{p \in \mathcal{P}} N_p. \tag{32}$$

Lemma 1, (31) and (32) together imply the following characterisation result for a subset of the Planck laws. The comments near the end of the Proof of Lemma 3 imply a similar result for the case $\alpha = \frac{1}{2}$.

Theorem 14 *Let $\lambda, r > 0$, $V \sim \text{beta}(\lambda, r)$, and $\alpha = c = 1$. Any two of the following implies the third:*

- (a) $X \stackrel{d}{=} \mathcal{V} \mathcal{L}_r X$ and $E(X) = (\lambda/a)(\zeta(\lambda + 1)/\zeta(\lambda))$;
- (b) $\mathcal{V} \stackrel{d}{=} V \prod_{p \in \mathcal{P}} U_{p,r}^{-1}$; and
- (c) $X \sim GP(\lambda; a, a, 1)$.

Theorem 14 does generalise Lemma 1, but at the expense of a complicated random scaling law and restrictions on the parameters. Another approach is to transfer complexity from the scaling to the weight function. Choose $w(x) \geq 0$ for $x \geq 0$ such that $0 < m_w := E(w(X)) < \infty$. Define the weighted distribution function $\mathcal{L}_w G = \int_0^x w(y) dG(y)/m_w$ and, if X has the distribution function G , then denote by $\mathcal{L}_w X$ a random variable having the distribution function $\mathcal{L}_w G$. It is quite easy to show that if $q \geq \lambda > 0$, $\tilde{V} \sim \text{beta}(q, 1)$ and $w(x) = q - \lambda + x$, then $X \stackrel{d}{=} \tilde{V} \mathcal{L}_w X$ if and only if $X \sim \text{gamma}(\lambda)$. See Pakes and Navarro (2007) for this and similar results. The following result (proved in §10) extends this to a larger family of Planck laws. The restrictions on α and c ensure that the weight function is non-negative.

Theorem 15 *Let either $\alpha \geq 1$ and $-\infty < c \leq 1$, or $0 < \alpha < 1$ and $-\alpha/(1 - \alpha) \leq c \leq 1$, and $\lambda > 0$ if $c < 1$ or $\lambda > 1$ if $c = 1$. If $q \geq \lambda$ and*

$$w(x) = q - \lambda + ax + \frac{cbxe^{-bx}}{1 - ce^{-bx}},$$

then any two of the following implies the third:

1.
 - (a) $X \stackrel{d}{=} \tilde{V} \mathcal{L}_w X$ and $m_w = q$;
 - (b) $\tilde{V} \sim \text{beta}(q, 1)$; and
 - (c) $X \sim GP(\lambda; a, b, c)$.

Theorems 1 and 9 can be interpreted in terms of weighting and scaling. The pdf of $Y := \gamma(\lambda)/a$ can be regarded as arising from weighting the $GP(\lambda; a, b, c)$ pdf; $Y \stackrel{d}{=} \mathcal{L}_w X$ where $w(x) = 1 - ce^{-bx}$. So if $W = \alpha/N$ then the representation in Theorem 1 can be expressed as

$$X \stackrel{d}{=} W \mathcal{L}_w X. \tag{33}$$

Since $P(W \leq 1) = 1$ it follows that X is stochastically smaller than Y , an observation which also is a consequence of the fact that the weight function is strictly increasing. Similarly, if $\nu > \lambda$ then the representation of $L(X)$ in Theorem 9 can be expressed as (33) where $w(x) = (1 - ce^{-bx})x^{\nu-\lambda}$ and $W = a/V$.

8 A fixed-point generalisation

We extend the weighting operation defined in the previous section by allowing X to have an arbitrary distribution function $G(x)$ supported in the real line and $w(x) \geq 0$ to be a weight function satisfying $0 < m_w := E[w(X)] < \infty$. We write $\mathcal{L}_w X_w$ to denote a random variable which has the weighted distribution function $\mathcal{L}_w G(x) := m_w^{-1} \int_{-\infty}^x w(v)dG(v)$. Finally let $H(x)$ be an arbitrary distribution function and let $0 < \beta < 1$ be a constant.

Here we are interested in fixed points of the transformation $\mathcal{T}G = (1 - \beta)\mathcal{L}_w G_w + \beta H$, that is, solutions for G of

$$dG(x) = cw(x)dG(x) + \beta dH(x), \tag{34}$$

where $c = (1 - \beta)/m_w$, thus ensuring that $G(\infty) = 1$. A formal solution of (34) is

$$\tilde{G}(x) := \beta \int_{-\infty}^x \frac{dH(v)}{1 - cw(v)}. \tag{35}$$

If $H(0-) = 0$ then $\tilde{G}(0-) = 0$ and $(1 - c)\tilde{G}(0) = \beta H(0)$, and in particular, if $c = 1$, then $H(0) = 0$ is a necessary condition for existence of a solution.

Relation (34) arises as the weak limit of the ‘in law’ recursion

$$X(n + 1) \stackrel{d}{=} (1 - I(n))\mathcal{L}_w X_w(n) + I(n)V(n), \quad (n = 0, 1, \dots) \tag{36}$$

where $X(0)$ has a specified distribution function, $I(n) \sim \text{Bern}(\beta)$, the distribution function of $V(n)$ is H , and $I(n)$, $\mathcal{L}_w X_w(n)$ and $V(n)$ are independent. The distribution function version of this recursion was analysed by Kingman (1978) in the particular case where all distribution functions are supported in $[0, 1]$ and $w(x) = x$. This case is his ‘house of cards’ model for genetic mutation-selection balance where $X(n)$ represents the fitness of an individual chosen randomly from the n th generation of an evolving population, the weight function represents the skewing effect of greater reproductive success of fitter individuals, β is a probability of mutation, and H expresses the fitness distribution of mutants.

The following result is a uniqueness assertion implying that \tilde{G} is determined by (34) for a given H .

Theorem 16 *Suppose that $0 \leq w(x) \leq 1$ and $\lim_{x \rightarrow 0} w(x) = 1$. Then (34) has a solution if and only if $0 < c \leq 1$, in which case (35) is the unique solution.*

The following theorem establishes a connection with the GP laws by representing the solution (35) in the spirit of Theorem 1.

Theorem 17 (i) *Let the conditions of Theorem 16 hold. Then*

$$\tilde{G}(x) = E[(\mathcal{W}_N H)(x)]$$

where $(\mathcal{W}_n H)(x) = \int_{-\infty}^x w^n(v) dH(v) / m_w(n)$, $m_w(n) = E[w^n(V(1))]$ and the random variable \mathcal{N} has the law

$$p_n := P(\mathcal{N} = n) = \beta c^n m_w(n), \quad (n = 0, 1, \dots). \tag{37}$$

(ii) Let $H(0-) = 0$, $\eta(s) := E(e^{-sV(1)})$ and $w(x) = e^{-bx}$ where b is a positive constant. Then

$$\tilde{G}(x) = E[H(x; b\mathcal{N})]$$

where $H(x; s) := \int_0^x e^{-sv} dH(v) / \eta(s)$ is a member of the natural exponential family generated by H , and

$$p_n = \beta c^n \eta(bn).$$

Proof Expand the denominator in (35) to obtain

$$\tilde{G}(x) = \beta \sum_{n=0}^{\infty} c^n \int_{-\infty}^x w^n(v) dH(v)$$

and interpret the summands in terms of quantities defined in the assertion. □

The transformation $Y = w(X)$ reduces the distribution function version of (36) to the recursion studied by Kingman (1978). The case $c < 1$ corresponds to his concept of ‘democracy’, and then the distribution functions $G(x; n) := P(X(n) \leq x)$ converge in total variation norm to $\tilde{G}(x)$ as given by (35). Moreover this convergence occurs geometrically fast. The case $c = 1$ is more problematic. The fact that \tilde{G} is a distribution function translates into the equality case of Kingman’s inequality (4.1). Next, the discrete probability distribution (37) is non-increasing and hence $\limsup_{n \rightarrow \infty} p_{n+1} / p_n \leq 1$. This condition is sufficient to ensure that the renewal sequence $\{u_n\}$ generated by the discrete law $\{p_n\}$ satisfies the strong ratio limit property $\lim_{n \rightarrow \infty} u_{n+1} / u_n = 1$; see Garsia et al. (1962). Thus Kingman’s meritocracy condition is fulfilled and hence we again have convergence to \tilde{G} in total variation norm. As a final remark, observe that the sequence $\{p_n\}$ is a Hausdorff moment sequence, and hence so is the renewal sequence $\{u_n\}$. Kingman (1978) proves this, and so too does Horn (1970). It follows that $\{u_n\}$ is a Kaluza sequence (Kingman (1972, §1.5)), that is, $\{u_{n+1} / u_n\}$ is non-decreasing.

It is evident now that the $GP(\lambda; a, b, c)$ law arises as the unique solution of (34) with the exponential weight in Theorem 17(ii), $V(1) = \gamma(\lambda) / a$ and $m_w = \beta \alpha^\lambda \Phi(c, \lambda, \alpha)$. We see too that the random scaling in Theorem 1 arises from the fact that members of the natural exponential family generated by a gamma law are simply scale changes of the generator. Finally, we observe that the case $\lambda = \frac{1}{2}$ is implicit in Turelli (, §3) where $V(1)$ has a centred normal law and the weight function has a Gaussian profile. The square of Turelli’s limiting fitness random variable has a generalized Planck law with $\lambda = \frac{1}{2}$; see his equation (3.6).

9 Related infinitely divisible laws

In this section we present some discrete and continuous infdiv laws which are related to the GP law in various ways.

1. Our first relation can be regarded as an alternative approach to the discussion of discrete Pareto laws in Steutel and van Harn (2004, p. 420). Recall that N_L has the Lerch law, $P(N_L = n) = p(n)$ where $\{p(n) : n = 0, 1, \dots\}$ is specified by the right-hand side of (9).

Also, let $(\mathcal{N}(\cdot))$ denote a unit rate Poisson process. The following result adds detail to Theorem 11. Its proof is in §10.

Theorem 18 *The Lerch law $L(N_L)$ is a Poisson mixture,*

$$p(n) = P(\mathcal{N}(T) = n), \quad (n = 0, 1, \dots)$$

where T is independent of \mathcal{N} and its density function is the mixture of exponential densities

$$f_T(x) = (cK)^{-1} e^x \int_1^\infty (\log v)^{\lambda-1} v^{-\alpha} e^{-vx/c} dv,$$

i.e., $T \stackrel{d}{=} c\varepsilon \exp(-\gamma(\lambda)/\alpha)$, where ε has a standard exponential law and it is independent of the gamma random variable $\gamma(\lambda)$. Hence $L(T)$ is infdiv.

2. Recall that $\mathcal{N}(\cdot)$ denotes the unit-rate Poisson process and that a positive random variable X has an infdiv law if and only if $\Delta := \mathcal{N}(\xi X)$ is infdiv (i.e., compound Poisson) for all $\xi > 0$ (Steutel and van Harn (2004, p. 369)). In view of the incompletely resolved infdiv status of the GP laws, it is of interest to see what emerges if $X \sim GP(\lambda; \alpha, c)$. In this case

$$\begin{aligned} p(j; \xi) := P(\Delta = j) &= \frac{\xi^j}{Kj!} \int_0^\infty \frac{x^{\lambda+j-1} e^{-(\alpha+\xi)x}}{1 - ce^{-x}} dx \\ &= \frac{\xi^j \Gamma(\lambda + j)}{\Gamma(\lambda)j!} \cdot \frac{\Phi(c, \lambda + j, \alpha + \xi)}{\Phi(c, \lambda, \alpha)}, \quad (j = 0, 1, \dots). \end{aligned}$$

It follows from Theorem 9 above that if $0 < c \leq 1$, then for any $v \geq \lambda$ this is a mixture of negative-binomial laws with common shape parameter v . In addition, it follows from Theorem 10 above and Theorem 12 in Steutel and van Harn (2004, p. 369) that $L(\Delta)$ is compound Poisson if $\lambda \leq 2$, and that $\{p(j; \xi)\}$ is log-convex if and only if $\lambda \leq 1$; see Theorem 6.13 in Steutel and van Harn (2004, p.374).

It is easy to see that the sequence $\{\xi^j \Gamma(\lambda + j)/j!\}$ is log-convex if and only if $\lambda \leq 1$, so the following result is a little unexpected.

Lemma 6 *Let $0 < c < 1$ and $\lambda > 0$, or $c = 1$ and $\lambda > 1$. The sequence defined by $\phi(j) = \Phi(c, \lambda + j, \alpha)$ for $j = 0, 1, \dots$ is log-convex. In addition the constant $\bar{K} := \sum_{j=0}^\infty \phi(j)$ is finite if and only if $\alpha > 1$, in which case*

$$\bar{K} = \sum_{n=0}^\infty \frac{c^n}{(\alpha - 1 + n)(\alpha + n)^{\lambda-1}}. \tag{38}$$

Proof Using the formal identity $\sum_{m \neq n} s_m \sigma_n = \sum_{m < n} (s_m \sigma_n + s_n \sigma_m)$, where $\{s_n\}$ and $\{\sigma_n\}$ are sequences, we have for $j \geq 1$ that

$$\phi(j-1)\phi(j+1) = \sum_{n \geq 0} \left(\frac{c^n}{(\alpha + n)\lambda} \right)^2 + \sum_{0 \leq m < n} [(\alpha + m)(\alpha + n)]^{-(\lambda+j-1)} [(\alpha + m)^{-2} + (\alpha + n)^{-2}].$$

The general inequality $u^2 + v^2 \geq 2uv$ implies that $\phi(j-1)\phi(j+1) \geq \phi^2(j)$, thus proving the first assertion. Next, (4) yields $\bar{K} = \sum_{n \geq 0} c^n \sum_{j \geq 0} (\alpha + n)^{-\lambda-j}$ and the inner sum is finite if and only if $\alpha > 1$. The inner sum is a geometric series, whence (38). \square

We thus obtain a compound Poisson law by normalising, i.e., define the discrete law

$$q(j) = \bar{K}^{-1} \Phi(c, \lambda + j, \alpha), \quad (j = 0, 1, \dots).$$

We remark in passing that the PGF $Q(s)$ of this law can be expressed as a kind of quasi-mixture with respect to $X \sim GP(\lambda; \alpha, c)$:

$$Q(s) = \frac{E[E_\lambda(sX)]}{E[E_\lambda(X)]},$$

where

$$E_\lambda(x) = \Gamma(\lambda) \sum_{j \geq 0} \frac{x^j}{\Gamma(\lambda + j)}, \quad (x \geq 0), \tag{39}$$

is the Mittag-Leffler function, i.e., a sub-family of confluent hypergeometric functions which generalises the exponential function; $E_1(x) = e^x$. See Erdelyi (1955, p. 210) or Olver et al. (2010, p. 261).

3. Let Y be a non-negative random variable with finite first order moment m_Y , and in this subsection only we denote the ordinary size-biased version of Y by $\hat{Y} := \mathcal{L}_1 Y$. Then the law of \hat{Y} is that of the stationary total lifetime in the renewal process generated by $L(Y)$. It is known that the total lifetime can be represented as the sum of Y and an independent increment if and only if $L(Y)$ is infdiv, that is, $L(Y)$ is infdiv if and only if

$$\hat{Y} \stackrel{d}{=} Y + Z \tag{40}$$

where $Z \geq 0$ and $Y \perp Z$. This is shown with the renewal context by van Harn and Steutel (1995), and independently by Pakes et al. (1996) emphasising the length-bias connection. Suppose that (40) holds and let $\psi(\theta)$ denote the cumulant function of $L(Y)$,

$$\psi(\theta) := -\log E(e^{-\theta Y}) = l_Y \theta + \int_0^\infty (1 - e^{-\theta x}) \ell(dx)$$

where l_Y is the infimum of the support of $L(Y)$ and $\ell(dx)$ is its Lévy measure, that is, a measure on \mathbf{R}_+ satisfying $\int_0^\infty (x \wedge 1) \ell(dx) < \infty$. The Laplace-Stieltjes transform of $L(Z)$ is $\xi(\theta) = \psi'(\theta)/m_Y$.

The following two examples explore the cases where one of $L(Y)$ or $L(Z)$ has a GP law.

Example 5 If $Y(\lambda) \sim GP(\lambda; a, b, c)$ then clearly $\hat{Y}(\lambda) \stackrel{d}{=} Y(\lambda + 1)$ for any $\lambda > 0$. So if $0 \leq c \leq 1$ and $0 < \lambda \leq 2$, then it follows from Theorem 10(i) and (40) that

$$Y(\lambda + 1) \stackrel{d}{=} Y(\lambda) + Z(\lambda) \quad \& \quad Y(\lambda) \perp Z(\lambda),$$

where, from (22), the Laplace-Stieltjes transform of $L(Z(\lambda))$ is

$$\xi(\theta; \lambda) = \frac{\Phi(c, \lambda, \alpha)}{\Phi(c, \lambda + 1, \alpha)} \cdot \frac{\Phi(c, \lambda + 1, \alpha + \theta/b)}{\Phi(c, \lambda, \alpha + \theta/b)}.$$

This emphasises that the question of whether the $GP(\lambda; a, b, c)$ law is infdiv is equivalent to deciding for which values of λ is $\xi(\theta; \lambda)$ a completely monotone function of θ .

In preparation for our second example we recall the following facts. Pakes et al. (1996) observe that if $Z = \gamma(\lambda)/\alpha$ in (40) then $L(Y)$ has the Hougaard law $H(\eta, \delta, \alpha)$, an infdiv law whose cumulant function is

$$\psi(\theta) = \frac{\delta}{\eta} [(\alpha + \theta)^\eta - \alpha^\eta],$$

where $\eta = 1 - \lambda < 1$ and $m_Y = \delta\alpha^{-\lambda}$. If $0 < \eta < 1$ then the Hougaard laws for $\alpha > 0$ comprise the natural exponential family generated by the positive stable law with index η ; see Seshadri (1993). The case $\eta = 0$ is just the gamma law $Y = \gamma(\delta)/\alpha$, and if $\eta < 0$ then $L(Y)$ is a compound Poisson law which has a gamma jump law.

Suppose in (40) that the distribution function of Z has the mixture form

$$G(x) = \sum_n p(n)G_n(x) \tag{41}$$

where $\{p(n)\}$ is an arbitrary discrete law and G_n is a distribution function satisfying $G_n(0) = 0$. Theorem 2.2 in Pakes et al. (1996) asserts that Y in (40) can be expressed as

$$Y \stackrel{d}{=} \sum_n Y_n \tag{42}$$

where the summands are independent and $L(Y_n)$ is infdiv with cumulant function

$$\psi_n(\theta) = m_Y p(n) \int_0^\infty (1 - e^{-\theta x}) x^{-1} dG_n(x).$$

Example 6 Choosing $Z \sim GP(\lambda; \alpha, c)$ in (40), then Theorem 1 implies the form (41) where the $p(n)$ are given by (9) and $G_n(x)$ is the distribution function of $\gamma(\lambda)/(\alpha + n)$. It follows that (42) holds with $Y_n \sim H(\eta, \delta_n, \alpha + n)$ where

$$\delta_n = \frac{m_Y c^n}{\Phi(c, \lambda, \alpha)},$$

and the cumulant function of Y can be shown to be

$$\psi(\theta) = m_Y \int_0^\infty (1 - e^{-\theta x}) x^{-1} g(x; \lambda, \alpha, c) dx.$$

It follows from the form of the Lévy density in this representation that $L(Y)$ is a compound Poisson law if $0 < c < 1$ and $\lambda > 1$, or $c = 1$ and $\lambda > 2$, and in either case the jump law is the $GP(\lambda - 1; \alpha, c)$ law; if $0 < c \leq 1$ and $1 < \lambda \leq 2$, then $L(Y)$ belongs to the Bondesson class of laws (i.e., the smallest class containing exponential mixtures and closed under convolution and weak limits); and if $0 < \lambda < 1$ and $0 \leq c \leq 1$, or if $\lambda = 1$ and $0 \leq c < 1$, then $L(Y)$ is a GGC law.

10 Proofs

Proof of Theorem 1 The change of variable $y = \alpha v$ in (5) yields

$$\Phi(c, \lambda, \alpha) = \frac{\alpha^{-\lambda}}{\Gamma(\lambda)} \int_0^\infty \frac{y^{\lambda-1} e^{-y}}{1 - ce^{-y/\alpha}} dy = \alpha^{-\lambda} E \left[\left(1 - ce^{-\gamma(\lambda)/\alpha} \right)^{-1} \right]. \tag{43}$$

It follows that the Laplace-Stieltjes transform of $(aX(\lambda) - \lambda)/\sqrt{\lambda}$ equals

$$\begin{aligned} e^{\theta\sqrt{\lambda}} \frac{\Phi(c, \lambda, a(1 + \theta/\sqrt{\lambda}))}{\Phi(c, \lambda, \alpha)} &= e^{\theta\sqrt{\lambda}} \left(1 + \theta/\sqrt{\lambda} \right)^{-\lambda} \\ &\times \frac{E \left[\left\{ 1 - c \exp \left(-\gamma(\lambda)/\alpha \left(1 + \theta/\sqrt{\lambda} \right) \right) \right\}^{-1} \right]}{E \left[\left\{ 1 - c \exp \left(-\gamma(\lambda)/\alpha \right) \right\}^{-1} \right]}. \end{aligned}$$

The exponents in the quotient of expectations converge to infinity in probability and hence this quotient converges to unity. The first factor on the right-hand side converges to $\exp(\frac{1}{2}\theta^2)$ since $(\gamma(\lambda) - \lambda) / \sqrt{\lambda} \xrightarrow{d} N(0, 1)$.

Proof of Theorem 4 Recall that $g \in SU$ if and only if (14) holds. If $\lambda < 1$, i.e., $\ell < 0$, then $C(x) > 0$ and Assertion (i) follows.

Suppose that $c < 0$. Then $C(x) < 0$ if $\lambda \geq 1$. If, instead, $0 < \lambda < 1$, then $\ell < 0$ and then (14) can be recast as

$$\frac{-\ell}{x^2} \leq \frac{(-c)e^x}{(e^x - c)^2}, \quad (x > 0).$$

This clearly cannot hold because, as $x \uparrow \infty$, the right-hand side decreases to zero much faster than the left-hand side. Assertion (ii) follows.

If $c > 0$ and $\lambda \geq 1$, i.e. $\ell \geq 0$, then $C(x)$ can be factorised as

$$C(x) = \left(\frac{\sqrt{c}e^{x/2}}{e^x - c} + \frac{\sqrt{\ell}}{x} \right) \left(\frac{\sqrt{c}e^{x/2}}{e^x - c} - \frac{\sqrt{\ell}}{x} \right). \tag{44}$$

The first factor is positive, so (14) holds if and only if the second factor is negative-valued, i.e.,

$$\frac{1}{2}x\sqrt{c} \leq \sqrt{\ell} \left(e^{\frac{1}{2}x} - ce^{-\frac{1}{2}x} \right), \quad (x > 0). \tag{45}$$

If $c = 1$, then this inequality is equivalent to $y \leq \sqrt{\ell} \sinh y$, where $y = x/2$. This inequality holds if and only if $\ell \geq 1$, i.e., $\lambda \geq 2$. Assertion (iii) is thus established.

Finally, let $0 < c < 1$. Obviously (14) cannot hold if $\lambda = 1$. So assume that $\lambda > 1$ and define the positive number $\varrho = \sqrt{c/\ell}$. Then, still with $y = x/2$, (45) becomes

$$2\varrho y \leq e^y - ce^{-y}. \tag{46}$$

The right-hand side is convex increasing from $1 - c > 0$ at $y = 0$. The parameter ϱ decreases through $(0, \infty)$ as ℓ increases from zero. Thus (46) holds if ℓ is sufficiently large. On the other hand, equality in (46) holds for two distinct values of y if ℓ is small. It follows that there is a critical value $\bar{\varrho} \in (0, \infty)$ such that the equality has a unique positive solution Υ for y . This configuration is equivalent to the conditions

$$2\varrho y = e^y - ce^{-y} \quad \text{and} \quad 2\varrho = e^y + ce^{-y};$$

the second condition is a tangent condition.

Eliminating ϱ yields

$$y = r(y) := \frac{e^{2y} - c}{e^{2y} + c}. \tag{47}$$

Observing that $r(0) = \varepsilon \in (0, 1)$, that $r(y) \uparrow 1$ as $y \uparrow \infty$, and that $r(y)$ is concave increasing, it follows that the unique solution $\Upsilon \in (\varepsilon, 1)$. Hence the critical value of r is

$$\varrho = \frac{1}{2} (e^\Upsilon + ce^{-\Upsilon}) = \sqrt{\frac{c}{1 - \Upsilon^2}}.$$

Consequently, (46) holds if and only if $r \leq \varrho$ and (47) can be re-arranged into the form (15). Assertion (iv) now follows.

Proof of Theorem 5 If $\alpha = c = 1$, then the sign equation (17) reduces to

$$\ell - y = \ell e^{-y}. \tag{48}$$

This has a solution $y = 0$ and it is the unique solution if and only if $\ell \leq 1$, i.e., $1 < \lambda \leq 2$. Assertion (i) follows.

If $\lambda > 2$, then $g(0) = 0$ and (48) has a unique positive solution in $(0, \ell)$, implying that g is unimodal. The value of the mode is obtained by rewriting (48) as $(y - \ell)e^{y-\ell} = -\ell e^{-\ell}$.

Observe that the right-hand side lies in $(-e^{-1}, 0)$ and recall that the functional equation $we^w = -z$, where $-e^{-1} < z < 0$, has two real-valued solutions, the principal and secondary Lambert functions of $-z$. The solution \bar{y} relevant for the present context tends to zero as $\ell \rightarrow 0$, and hence $-(\bar{y} - \ell) = W(-\ell e^{-\ell})$, where $W(\cdot)$ denotes the principal Lambert function. Assertion (ii) now follows because $M_O = a\bar{y}$. Assertion (iii) follows from the expansion $W(x) = x + o(x)$ as $x \rightarrow 0$.

Proof of Theorem 6 The relation (17) with $c = 1$ takes the form

$$\ell - \alpha y = R(y) := \frac{y}{e^y - 1}. \tag{49}$$

Clearly $R(0) = 1$ and $R(y)$ decreases convexly to zero as $y \uparrow \infty$. Assertion (i) follows from graphical considerations. Assertion (ii) follows similarly by considering intersections of the graphs of each side of (49). In particular, for a fixed ℓ the critical value α_c corresponds to a solution of (49) subject to the tangent condition $-\alpha = R'(y)$. Eliminating α yields the equation $z^{-1} \sinh z = \ell^{-\frac{1}{2}}$, where $z = \frac{1}{2}y$. Denoting the solution by \bar{z} , (49) yields the critical value as

$$\alpha_c = \frac{\ell}{2\bar{z}} - \frac{1}{e^{2\bar{z}} - 1}.$$

Observing that $R'(0) = -\frac{1}{2}$, it follows that $\alpha_c < \frac{1}{2}$. The reasoning is similar if $\ell = 1$ and the assertion follows because the two graphs intersect at $y = 0$.

Proof of Theorem 7 Considering only the case $\lambda > 1$, observe that $S(0) = (\lambda - 1)(1 - c) > 0$ and hence there is at least one positive mode. Next, we have

$$S''(y) = (\ell - 2\alpha - \alpha y)e^y,$$

independent of c . It follows that $S(y)$ is concave in $(0, \infty)$ if $2\alpha \geq \ell$, i.e. if $a/b \geq \frac{1}{2}(\lambda - 1)$, and Assertion (iii) follows.

If $2\alpha < \ell$, then $S(\cdot)$ has a unique inflection point $y_i = (\ell/\alpha) - 2 \in (0, \infty)$ and it is convex in $[0, y_i]$ and concave in $[y_i, \infty)$.

Now

$$S'(0) = \ell - \alpha(1 - c) - c \begin{cases} > 0 & \text{if } \ell > \bar{\ell} := \alpha(1 - c) + c \\ < 0 & \text{if } \ell < \bar{\ell}. \end{cases}$$

So if $\ell > \bar{\ell}$, then $S(\cdot)$ is convex increasing in $[0, y_i]$. It follows that $S(\cdot)$ has a unique global maximum, at y_M say, and a unique zero y_0 with $y_i < y_M < y_0 < \ell/\bar{\ell}$. Hence $g(x)$ is unimodal.

If $\ell < \bar{\ell}$, then $S(\cdot)$ is convex decreasing near $y = 0$. The above assertions about y_M and y_0 hold (though y_0 may not be the only zero), and now $S(\cdot)$ has a local minimum $y_m \in (0, y_i)$. If $S(y_m) < 0$, then there are a further two zeros y_{i0} such that $0 < y_{i0} < y_m < y_{20} < y_M$ and $g(x)$ is bimodal in such a case.

Proof of Theorem 8 Write $k = -c$ and observe that (49) becomes

$$\alpha y - \ell = R(y) := \frac{ky}{e^y + k}. \tag{50}$$

Clearly $R(0) = 0$ and $R'(0) = k/(1 + k) > 0$. A calculation shows that R has a global maximum at $\bar{y} = 1 + W(ke^{-1}) > 1$, an inflection point $y_i \in (\bar{y}, \infty)$, and it is concave in

$(0, y_i)$ and convex in (y_i, ∞) . Assertion (i) follows from these observations and that $\ell > 0$ in this case.

If $\ell = 0$, then $g(0) \in (0, \infty)$, $g'(0) > 0$, and $y = 0$ solves (50) for any $\alpha > 0$. It is the unique solution if $\alpha \geq R'(0)$, and there is a second positive solution, M_0 say, if $\alpha < R'(0)$. Assertion (ii) now follows.

If $0 < \lambda < 1$, then the left-hand side of (50) has a positive intercept $-\ell$ at $y = 0$. Hence (50) has no positive solution for any $\alpha > 0$ if $-\ell \geq R(\bar{y}) = \bar{y} - 1 = W(ke^{-1})$, i.e., if $\lambda \geq \lambda_c := 1 + W(ke^{-1})$. Assertion (iii) and the first part of (iiia) follow. This outcome can occur only if $\ell \geq \max R(y) = R(\bar{y}) = W(ke^{-1})$, i.e., $\ell e^\ell \leq ke^{-1}$. The second part of Assertion (iiia) follows.

If $\lambda < \lambda_c$, then $\ell < R(\bar{y})$ and (50) has no solution if α is large and it has two positive solutions if $\alpha \approx 0$. Consequently, there exists a critical value of α such that (50) has a unique positive solution, implying that the graphs of $\alpha y - \ell$ and $R(y)$ meet tangentially. This unique solution y_t and the critical value α_c are obtained by solving (50) subject to $\alpha = R'(y)$. Eliminating α gives the equation $-\ell = R(y)(y - R(y))$. This has two solutions and y_t is the smaller. Thus $y_t < \bar{y}$ and $\alpha_c = y_t^{-1}(R(y_t) + \ell)$. Assertion (iiib) follows from these facts.

Proof of Theorem 10 Bondesson (1987) defines a class I_c of pdf's f on $(0, \infty)$ which are positive with a continuous derivative and satisfy

$$\frac{f'(y)}{f(y)} \leq \frac{1}{x} + \frac{f'(x)}{f(x)}, \quad (0 < y \leq x).$$

The class I_c properly contains log-convex pdf's. Members of I_c are decreasing and they are infdiv (Bondesson (1987, Theorem 4.1)). Finally, a pdf $f \in I_c$ has a representation

$$f(x) = f(1) \exp\left(-\int_1^x \left(\psi(y) + \frac{l(y)}{y}\right) dy\right), \quad (x > 0), \tag{51}$$

where ψ is continuous and positive-valued and l is continuous with range in $[0, 1]$. Conversely, any pdf having this form is in I_c (Bondesson (1987, Theorem 4.2)).

An equivalent expression of (51) is

$$-\frac{f'(x)}{f(x)} = \psi(x) + \frac{l(x)}{x}.$$

Referring to (16), let $k = -c$ (which is positive) to obtain

$$-\frac{g'(x)}{g(x)} = a + \frac{1 - \lambda}{x} - \frac{kb}{e^{bx} + k}.$$

Choose $\psi(x) \equiv a$ and $l(x) = 1 - \lambda - R(bx)$ where, recalling (50), $R(y) = ky/(e^y + k)$. Bondesson's Theorem 4.2 is applicable if $0 \leq l(x) \leq 1$. The right-hand side inequality obviously holds, and the left-hand side inequality is equivalent to the condition $1 - \lambda \geq \max_{y \geq 0} R(y)$.

It follows as in the Proof of Theorem 5 that R has a unique maximum at $y = \bar{y} = 1 + W(ke^{-1})$ and that $R(\bar{y}) = \bar{y} - 1 = W(ke^{-1})$. Hence $g \in I_c$ if and only if $1 - \lambda \geq W(ke^{-1})$. This inequality implies that $\lambda \leq 1$ and that $W(ke^{-1}) < 1 = W(e)$, i.e., $k < e^2$. Assertion (ii) follows.

Remark 8 The conditions (20) appear to be optimal for the conclusion. Specifying an arbitrary non-negative continuous decreasing function ψ , Bondesson's representation

requires that $\psi(x) \downarrow a$. This leads to a form of $l(x)$ uniformly smaller than the above choice and hence to more stringent conditions for it to be non-negative.

Proof of Proposition 1 Bondesson (1987) defines the class $S_c \subset I_c$ of pdf's f on $(0, \infty)$ which are positive-valued, $f''(x)$ is continuous and which satisfy

$$\frac{f'(y)}{f(y)} \leq \frac{1}{x} + \frac{f''(x)}{f'(x)}, \quad 0 < y \leq x < \infty.$$

Members of S_c are SD.

Applied to the GP pdf $g(x)$ and recalling that $\gamma(x) = g'(x)/g(x)$, Bondesson's criterion becomes

$$\gamma(y) \leq \frac{1}{x} + \frac{\gamma'(x)}{\gamma(x)} + \gamma(x), \quad 0 < y \leq x < \infty. \tag{52}$$

Now $\gamma(0+) = -\infty$ and $\gamma(\infty-) = -a$, so if $\gamma'(x) \geq 0$ in $(0, \infty)$, i.e., g is log-convex, then $\gamma(x) < 0$ in $(0, \infty)$ and (52) simplifies to

$$\gamma(x) + x\gamma'(x) \leq 0, \quad (x > 0).$$

This condition takes the explicit form (where $y = bx$),

$$D(y) := \frac{cye^y}{(e^y - c)^2} - \frac{c}{e^y - c} \leq \alpha = a/b, \quad (y > 0),$$

i.e.,

$$\max_{y>0} D(y) \leq \alpha. \tag{53}$$

If $0 < c \leq 1$, then $\gamma'(x) \geq 0$, so we have only to check (53). In the case $c = 1$ we have

$$D(y) = \frac{ye^y - e^y + 1}{(e^y - 1)^2},$$

and $D(0+) = \frac{1}{2}$. In addition

$$D'(y) = \frac{e^y}{(e^y - 1)^3} [(y + 2)(e^y - 1) - 2ye^y].$$

The square bracket term equals

$$-\sum_{n \geq 2} \left(\frac{1}{(n-1)!} - \frac{2}{n!} \right) y^n < 0.$$

Hence $D(y)$ is decreasing, so $\max_{y>0} D(y) = \frac{1}{2}$, and Assertion (i) follows.

If $0 < c < 1$, then the relation $D'(y) = 0$ can be cast as

$$e^y = c\phi(y) \quad \text{where} \quad \phi(y) = \frac{2+y}{2-y}, \quad (y > 0, n \in \mathbb{N}). \tag{54}$$

Clearly ϕ is positive-valued and increasing in $(0, 2)$, $\phi(2-) = \infty$, and ϕ is negative-valued and decreasing in $(2, \infty)$. Hence there is a single value y_c of y satisfying (54) and $y_c \in (0, 2)$. Moreover, since $D'(0) = c(1-c)^{-2} > 0$,

$$\max_{y>0} D(y) = D(y_c) = \frac{1}{2} \left(1 - \frac{1}{2} y_c \right).$$

Assertion (ii) follows now from (53).

Assume now that $c < 0$ and write $k = -c$. Consider first the sign of $\gamma'(x)$, equivalently, the sign of

$$\psi(y) := \frac{L}{y^2} - \frac{ke^y}{(e^y + k)^2},$$

where $L = 1 - \lambda \in [0, 1)$. We have $\psi(y) \geq 0$ in $(0, \infty)$ if and only if

$$\max_{y>0} \xi(y) \leq \sqrt{\frac{L}{k}}, \quad \text{where } \xi(y) = \frac{ye^{y/2}}{e^y + k}.$$

Critical points of ξ solve the equation

$$A(y) = \frac{1}{y} - \frac{e^y}{e^y + k} = -\frac{1}{2}. \tag{55}$$

Clearly $A(0+) = \infty$, $A(\infty-) = -1$ and A is decreasing. Hence there is exactly one critical point y_c and it is a global maximum of ξ because $\xi(0) = \xi(\infty-) = 0$. Now (55) rearranges into the form

$$e^y = k \frac{y+2}{y-2};$$

c.f. (23).

It follows, as above, that $y_c > 2$ and that

$$\xi(y_c) = \sqrt{\frac{y_c^2 - 4}{4k}} \leq \sqrt{\frac{1 - \lambda}{4k}}.$$

Hence $\gamma'(x) \geq 0$ if and only if $(y_c^2/4) - 1 \leq 1 - \lambda$, i.e., the bound in Assertion (iii) holds. Checking that the log-derivative of $e^y(y-2)/(y+2)$ is positive in $(2, \infty)$, it follows that y_c increases through $(2, \infty)$ as k increases through $(0, \infty)$. Hence the left-hand inequality in (23) can hold only if $y_c < 2\sqrt{2}$, i.e., only if

$$0 < k < e^{2\sqrt{2}} \frac{\sqrt{2}-1}{\sqrt{2}+1} = (3-2\sqrt{2})e^{2\sqrt{2}} = -c'.$$

(This establishes assertions in Remark 4.)

Next, in terms of k we see that

$$D(y) = \frac{k}{e^y + k} - \frac{kye^y}{(e^y + k)^2},$$

$D(0) = k/(k+1) \in (0, 1)$ and $D(y) \sim -kye^{-y} < 0$ as $y \rightarrow \infty$. As in the previous case, $D'(y) = 0$ if and only if (54) (with $c = -k$) holds. That equation has a unique solution $y_c \in (0, 2)$. But $D'(0) = -2k(k+1)^{-2} < 0$, hence $\min_{y>0} D(y) = D(y_c) < 0$. It follows that $\max_{y \geq 0} D(y) = D(0)$, and Assertion (iii) follows.

Proof of Lemma 3 Theorem 1 allows us to write $\log X = \Lambda - \mathcal{N}$ where $\Lambda = \log \gamma(\lambda)$ and $\mathcal{N} = \log N$. The assertion follows because Λ and \mathcal{N} have infdiv laws. To see this observe first that the moment generating function $\mu_\Lambda(t)$ of Λ is $\Gamma(\lambda + t)/\Gamma(\lambda)$. It follows that $E(\Lambda) = \psi(\lambda)$ where $\psi(t) = (d/dt) \log \Gamma(t)$ is the digamma function, and

$$\frac{d^2}{dt^2} \log \mu_\Lambda(t) = \psi'(\lambda + t) = \int_0^\infty \frac{v \exp(-v(\lambda + t))}{1 - \exp(-v)} dv;$$

see Gradshteyn and Ryzhik (1980, 8.361(1)). Integration yields the Kolmogorov representation

$$\log \mu_\Lambda(t) = \text{const.}t + \int_0^\infty (e^{vt} - 1 - vt) v^{-2} \kappa(dv)$$

where

$$\kappa(dv) = \frac{ve^{-\lambda v}}{1 - e^{-v}} dv, \quad (v > 0).$$

Hence $L(\Lambda)$ is a spectrally-positive infdiv law.

The law of \mathcal{N} in the case $b = a$, i.e. $\alpha = 1$, is the zeta law, and this was shown to be infdiv by Khintchine in 1938. See Steutel and van Harn (2004, p. 131) and, with more detail, Lin and Hu (2001). The essence of this is that the Euler product representation of the Riemann zeta function leads to a representation of \mathcal{N} as a weighted sum of independent random variables,

$$\mathcal{N} \stackrel{d}{=} \sum_{p \in \mathcal{P}} (\log p) W_p, \tag{56}$$

where \mathcal{P} is the set of prime numbers and $P(W_p = n) = (1 - p^{-\lambda})p^{-\lambda n}$. Hence \mathcal{N} is a weighted sum of independent random variables having geometric laws, and since these are infdiv, then $L(\mathcal{N})$ is infdiv too. Its Lévy measure $\ell_1(dx)$ assigns mass $(mp^{m\lambda})^{-1}$ to $m \log p$, $p \in \mathcal{P}$ and $m \in \mathbb{N}$. The case $\alpha = \frac{1}{2}$ is treated by Hu et al. (2006). They show in this case that

$$\mathcal{N} \stackrel{d}{=} \sum_{p \in \mathcal{P}, p > 2} (\log p) W_p \quad \& \quad \ell_{\frac{1}{2}}(dx) = \ell_1(dx) - \sum_{m=1}^{\infty} 2^{-m\lambda} \delta_2(dx).$$

Hu et al. (2006) prove the surprising result that $L(\mathcal{N})$ is infdiv only if $\alpha = \frac{1}{2}$ or $\alpha = 1$.

Proof of Theorem 12 Recall the definitions of GGC and HCM near the end of §6 and widen the definition of a GGC law by admitting any function as a GGC if it is proportional to the pdf of a GGC law. The proof requires the following concepts and facts.

- (a) If a probability mass function $p(n)$ is a GNBC and if $c > 0$ and $\sum_{n \geq 0} p(n)c^n < \infty$, then the probability mass function $\propto p(n)c^n$ is a GNBC, (Bondesson (1992, p. 128));
- (b) An absolutely monotone probability mass function

$$p(n) = \int_0^{\infty} e^{-nx} k(x) dx \tag{57}$$

can be expressed as the Poisson mixture

$$p(n) = \frac{1}{n!} \int_0^{\infty} z^n e^{-z} g(z) dz$$

where

$$g(z) := \int_0^{\infty} e^{-zy} \xi(y) dy$$

and $\xi(y) = k(\log(1 + y))$. The discrete law $\{p(n)\}$ is a GNBC if and only if g is a GGC pdf (Bondesson (1992, p. 127) or Steutel and van Harn (2004, p. 391));

- (c) $L(Y)$ is a GGC if and only if its Laplace-Stieltjes transform is HCM, (Bondesson (1992, p. 81) or Steutel and van Harn (2004, p. 362));
- (d) An HCM function is a GGC, (Bondesson (1992, p. 57) or Steutel and van Harn (2004, p. 358));
- (e) Positive powers and products of HCM functions are HCM;

- (f) If $\eta(x)$ is HCM then so is the composition $\eta(h(y))$ where $h(y) = \int_0^1 y^t \bar{m}(dt)$ and \bar{m} is a measure supported in $[0, 1]$, (a limiting form of Theorem 3.3.2 in Bondesson (1992, p. 41));
- (g) If $\kappa(\theta)$ is a decreasing HCM function then $\kappa(a + \theta)$ is HCM for any $a > 0$, (Bondesson (1992, p. 68)); and
- (h) The limit of a sequence of HCM functions is HCM.

The proof. Assume Case (i), that $\lambda > 1$ In view of (a) it suffices to assume $c = 1$. The case $\lambda = 1$ is obvious from what follows, so we assume $\lambda > 1$. Ignoring the normalisation constant $\Phi(1, \lambda, \alpha)$ in the definition (9), $p(n)$ has the form (57) where $k(x) \propto x^{\lambda-1} e^{-\alpha x}$, giving

$$\xi(y) = (1 + y)^{-\alpha} [\log(1 + y)]^{\lambda-1} = y^{\lambda-1} (1 + y)^{-\alpha} \left(\frac{\log(1 + y)}{y} \right)^{\lambda-1}.$$

It follows that $\int_0^\infty g(z) dz < \infty$. If we can show that $\xi(y)$ is HCM then it follows from (c) and (d) that $g(z)$ is HCM, and hence is a GGC. The assertion will follow then from (b).

It is obvious that the identity function and the function $(1 + y)^{-1}$ are HCM, so it follows from the form of $\xi(y)$ and from (e) that we only need to show that $y^{-1} \log(1 + y)$ is HCM, and this follows from (f) - (h) as follows.

Clearly $\eta(x) = (1 + \delta)/(\delta + x)$ is HCM for all $\delta > 0$. Choosing $\bar{m}(dt) = dt$ gives $h(y) = -(1 - y)/\log y$ if $y > 0$ and $y \neq 1$, and $h(1) = 1$, so we conclude from (f) that $\kappa(y) = \eta(h(y))$ is HCM. It is decreasing too, so it follows from (g) that

$$\kappa(1 + y)/\kappa(1) = \frac{(1 + \delta) \log(1 + y)}{y + \delta \log(1 + y)}$$

is HCM. The desired outcome now follows by letting $\delta \rightarrow 0$ and from (h).

Case (ii), $\alpha > 1$ and $\lambda > 0$, is the particular case $\beta = 1$ of Theorem 8.3.3 in Bondesson (1992). The essence of his proof is that $\xi(y)$ is proportional to the pdf of $Y = e^{(\alpha-1)y^\lambda} - 1$ and the Proof of Theorem 6.2.6 in Bondesson (1992) shows that $L(Y)$ is a GGC. Hence the Laplace transform $g(z)$ is HCM. However, $\int_0^\infty g(z) dz = \infty$, but it follows that g is the density function of a wide sense GGC (Bondesson (1992, p. 46)). This suffices for the conclusion.

Proof of Theorem 15 The restrictions on α and c ensure that $w(x) \geq 0$. If (a) and (b) both hold, then the distribution function G of X is absolutely continuous with a pdf $g(x) = G'(x)$ for almost every $x > 0$. In addition $G(0) = (\mathcal{L}_w G)(0) = G(0)w(0)/q$. But $w(0) = q - \lambda + \delta_{c,1}$, so $G(0) > 0$ if and only if $\lambda = \delta_{c,1}$, that is, if and only if $\lambda = c = 1$. This violates our assumptions. Hence we conclude that under any pair of the conditions (a) to (c), $G(x)$ has a pdf $g(x)$. Consequently the weighted distribution function has the pdf $w(x)g(x)/m_w$.

If (a) and (b) hold, then these pdf's are related by

$$g(x) = q^{-1} x^{q-1} \int_x^\infty w(y)g(y)y^{-q} dy.$$

This has a unique solution which can be confirmed to be (3) by observing that in this case

$$w(x) = q - 1 - xg'(x)/g(x), \tag{58}$$

and a little algebra shows that the differential equation derived from the integral equation indeed is satisfied.

This argument shows also that if (b) and (c) hold, then the first part of (a) holds, and we only need to observe that integrating (58) yields

$$m_w = q - 1 - \int_0^\infty xg'(x)dx = q,$$

after integrating by parts. Thus we have proved that (a) and (c) are equivalent if (b) holds.

Now assume that (a) and (c) hold, and let $H(y)$ denote the distribution function of $V(1)$. A Mellin convolution for the density function of a quotient of independent random variables yields

$$\begin{aligned} g(x) &= q^{-1} \int_0^\infty w(x/y)g(x/y)y^{-1}dH(y) \\ &= (1 - q^{-1}) \int_0^\infty g(x/y)y^{-1}dH(y) - q^{-1} \int_0^\infty (x/y)g'(x/y)y^{-1}dH(y), \end{aligned}$$

where we have used (58) for the second equality. Letting $N(t) = E[(V(1))^t]$, it follows that

$$\begin{aligned} M_g(t) &= (1 - q^{-1})N(t)M_g(t) - q^{-1}N(t) \int_0^\infty x^{t+1}g'(x)dx \\ &= (1 - q^{-1})N(t)M_g(t) + q^{-1}(t + 1)M_g(t)N(t). \end{aligned}$$

It follows that $N(t) = q/(q + t)$, and this is the moment function of the beta($q, 1$) law. Thus (a) and (c) imply (b).

Proof of Theorem 16 Suppose G solves (34). Clearly $G \ll H$, so

$$1 = cw(x) + \beta \frac{dH}{dG}(x) \geq cw(x).$$

Letting $x \rightarrow 0$ shows that $c \leq 1$. Suppose there are two distinct distribution function solutions and let $\Delta(x)$ be their difference. Then $\Delta(\infty) = 0$, and

$$\Delta(x) = c \int_{-\infty}^x w(v)d\Delta(v) = \dots = c^n \int_{-\infty}^x w^n(v)d\Delta(v). \tag{59}$$

Observing that the last integral is bounded above in magnitude by 2, allowing $n \rightarrow \infty$ implies that $\Delta(x) \equiv 0$ if $c < 1$. If $c = 1$ then $\Delta(x) = \int_{\mathcal{I}} d\Delta(x)$ where $\mathcal{I} = \{x : w(x) = 1\}$. This integral equals $\lim_{x \rightarrow \infty} \int_{-\infty}^x w^n(x)d\Delta(x) = \Delta(\infty) = 0$. Finally, \tilde{G} is a distribution function because

$$\int_{-\infty}^\infty d\tilde{G}(x) = \beta \int_{-\infty}^\infty \frac{cw(x) + (1 - cw(x))dH(x)}{1 - cw(x)} = cm_w + \beta = 1.$$

Proof of Theorem 18 Let $Z \sim GP(\lambda; \alpha, c)$. Then (3) and (6) yield

$$E \left[(1 - ce^{-Z})e^{-\theta Z} \right] = K^{-1} \int_0^\infty z^{\lambda-1} e^{-(\alpha+\theta)z} dz = (\Phi(c, \lambda, \alpha)(\alpha + \theta)^\lambda)^{-1}.$$

Comparing this with (9) gives the following representation as a mixture of geometric laws, $p(n) = E[(1 - U)U^n]$, where $U = ce^{-Z}$ and whose density function is

$$f_U(u) = \begin{cases} \frac{(\log(c/u))^{\lambda-1}(u/c)^\alpha}{K u(1-u)} & \text{if } 0 < u < c, \\ 0 & \text{otherwise.} \end{cases}$$

Now let ε have a standard exponential law and be independent of T and \mathcal{N} . Since

$$\frac{1 - u}{1 - us} = E \left(\exp(-(1 - s)\mathcal{N}(\varepsilon/(u^{-1} - 1))) \right),$$

the Poisson mixture form evidently holds with

$$T = \frac{\varepsilon}{U^{-1} - 1} = \frac{\varepsilon c e^{-Z}}{1 - c e^{-Z}},$$

and ε , Z and \mathcal{N} are mutually independent. This shows that $L(T)$ is a mixed exponential law, and hence it is *infdiv*. The geometric mixture representation implies that $\{p(n)\}$ is completely monotone and hence log convex (Steutel and van Harn (2004, p. 377)). This implies that $\{p(n)\}$ is *infdiv*.

Abbreviations

GP: Generalised Planck; pdf: Probability density function; *infdiv*: Infinitely divisible; CIR: Cox-Ingersoll-Ross; SD: Self-decomposable; BDLP: Background driving Lévy process; GGC: Generalised gamma convolution; GNBC: Generalised negative binomial convolution; HCM: Hyperbolically completely monotone

Acknowledgments

Remark 7 acknowledges a remark made by a referee of a much earlier version of my paper. I thank the referees of the present version for their care and remarks which have improved my presentation.

Authors' contributions

The author read and approved the final manuscript.

Funding

This research has no funding support.

Declarations

Availability of data and materials

Not applicable.

Competing interests

The author declares that he has no competing interests.

Received: 13 January 2021 Accepted: 15 June 2021

Published online: 01 August 2021

References

- Berg, C.: On powers of Stieltjes moment sequences. *I. J. Theor. Prob.* **18**, 871–889 (2005)
- Brigo, D., Mercurio, F.: *Interest Rate Models - Theory and Practice*. Springer, Berlin (2001)
- Bondesson, L.: On the infinite divisibility of the half-Cauchy and other decreasing densities and probability functions on the nonnegative half line. *Scand. Actuar. J.* **3-4**, 225–247 (1987)
- Bondesson, L.: *Generalized Gamma Convolutions and Related Classes of Distributions and Densities*. Springer-Verlag, New York (1992)
- Bondesson, L.: A class of probability distributions that is closed with respect to addition as well as multiplication of independent random variables. *J. Theor. Probab.* **28**, 1063–1081 (2015)
- Cox, J., Ingersoll, J., Ross, S.: A theory of the term structure of interest rates. *Econometrica*. **53**, 385–408 (1985)
- Cropper, W. H.: *The Quantum Physicists*. O.U.P., New York (1970)
- Davis, H. T.: *The Theory of Econometrics*. Principia Press, Bloomington (1941)
- Dennis, B., Patil, G. P.: The gamma distribution and weighted multimodal gamma distributions as models of population abundance. *Math. Biosci.* **68**, 187–212 (1984)
- Dharmadhikari, S., Joag-Dev, K.: *Unimodality, Convexity, and Applications*. Academic Press, New York (1998)
- Erdélyi, A., *Functions, Higher Transcendental*, Vol. III (Eds). McGraw-Hill, New York (1955)
- Feller, W.: Two singular diffusion problems. *Ann. Math.* **54**, 173–182 (1951)
- Garsia, A. M., Orey, S., Rodemich, E.: Asymptotic behaviour of successive coefficients of some power series. *Illinois J. Math.* **6**, 620–629 (1962)
- Gradshteyn, I. S., Ryzhik, I. M.: *Tables of Integrals, Series and Products*. Academic Press, New York (1980). The 8th edition is available online
- Gupta, P. L., Gupta, R. C., Ong, S.-H., Srivasta, H. M.: A class of Hurwitz-Lerch distributions and their applications in reliability. *Appl. Math. Comput.* **196**, 521–531 (2008)
- van Harn, K., Steutel, F.: Infinite divisibility and the waiting time paradox. *Stoch. Model.* **11**, 527–540 (1995)
- Horn, R. A.: On moment sequences and renewal sequences. *J. Math. Anal. Appl.* **31**, 130–135 (1970)
- Horsthemke, W., Lefever, R.: *Noise Induced Transitions*. Springer, Berlin (1984)
- Hu, C.-Y., Iksanov, A., Lin, G. D., Zakusylo, O. K.: The Hurwitz zeta distribution. *Austral. N.Z.J. Statist.* **48**, 1–6 (2006)
- Johnson, N. L., Kotz, S.: *Continuous Univariate Distributions - 2*. Houghton-Mifflin Co., Boston (1970)
- Johnson, N. L., Kemp, A. W., Kotz, S.: *Univariate Discrete Distributions*, 3rd ed. Wiley, New York (2005)
- Kingman, J. F. C.: *Regenerative Phenomena*. Wiley, London (1972)
- Kingman, J. F. C.: A simple model for the balance between selection and mutation. *J. Appl. Prob.* **15**, 1–12 (1978)
- Kittel, C.: *Thermal Physics*. Wiley, New York (1969)

- Klar, B., Parthasarathy, P. R., Henze, N.: Zipf and Lerch limit of birth and death processes. *Prob. Eng. Info. Sci.* **24**, 129–144 (2010)
- Kleiber, C., Kotz, S.: *Statistical Size Distributions in Economics and Actuarial Sciences*. Wiley-Interscience, New York (2003)
- Kotz, S., Kozubowski, T. J., Podgórski, K.: *The Laplace Distribution and Generalizations*. Birkhäuser, Boston (2001)
- Lamberton, D., Lapeyre, B.: *Introduction to Stochastic Calculus Applied to Finance*. Chapman-Hall, London (1996)
- Li, Z.: *Measure-Valued Branching Processes*. Springer, Berlin (2011)
- Lin, G. D., Hu, C.-Y.: The Riemann zeta distribution. *Bernoulli*. **7**, 817–828 (2001)
- Longair, M. S.: *Theoretical Concepts in Physics*. C.U.P., Cambridge (1984)
- Lukacs, E.: A characterization of the gamma distribution. *Ann. Math. Statist.* **26**, 319–324 (1965)
- Nadarajah, S., Kotz, S.: A generalized Planck distribution. *TEST*. **15**, 361–374 (2006)
- Olver, F. W., Lozier, D. W., Boisvert, R. F., Clark, C. W.: *NIST Handbook of Mathematical Functions*. C.U.P., New York (2010)
- Pais, A.: *Subtle is the Lord*. O.U.P., New York (1982)
- Pakes, A. G.: Length biasing and laws equivalent to the log-normal. *J. Math. Anal. Appl.* **197**, 825–854 (1996)
- Pakes, A. G.: Characterization by invariance under length-biasing and random scaling. *J. Statist. Planning Inf.* **63**, 285–310 (1997)
- Pakes, A. G., Navarro, J.: Distributional characterizations through scaling relations. *Aust. N.Z. J. Stat.* **49**, 115–135 (2007)
- Pakes, A. G., Sapatinas, T., Fosam, E. B.: Characterizations, length-biasing, and infinite divisibility. *Statist. Pap.* **37**, 53–69 (1996)
- Puertas-Centeno, D., Toranzo, I., Dehasa, J.: Biparametric complexities and generalized Planck radiation law. *J. Phys. A: Math. Theor.* **50**, 505001 (22pp.) (2017)
- Sato, K.-I.: *Lévy Processes and Infinitely Divisible Distributions*, Revised ed. C.U.P, Cambridge (2013)
- Schilling, R. L., Song, R., Vondraček, Z.: *Bernstein Functions*, 2nd. ed. De Gruyter, Berlin (2012)
- Seshadri, V.: *Inverse-Gaussian Distributions: A Case Study in Natural Exponential Families*. Clarendon Press, Oxford (1993)
- Stutel, F. W., van Harn, K.: *Infinite Divisibility of Probability Distributions on the Real Line*. Marcel Dekker Inc., New York (2004)
- Stewart, S. M.: Blackbody radiation functions and polylogarithms. *J. Quant. Spectrosc. Radiative Trans.* **113**, 232–238 (2012)
- Stuart, A.: Gamma-distributed products of independent random variables. *Biometrika*. **49**, 564–565 (1962)
- Tomovski, Ž., Saxena, R., Pogány, T.: Probability distributions associated with Mathieu type series. *ProbStat. Forum.* **05**, 86–96 (2012)
- Turelli, M.: Heritable genetic variation via mutation-selection balance: Lerch's zeta meets the abdominal bristle. *Theor. Popul. Biol.* **25**, 138–193 (1984)
- Varrò, S.: Irreducible decomposition of Gaussian distributions and the spectrum of black-body radiation. *Physica Scripta.* **75**, 160–169 (2007)
- Valluri, S., Gil, M., Jeffrey, J., Basu, S.: The Lambert W function and quantum statistics. *J. Math. Phys.* **50**, 102103 (2009)
- Weinberg, S.: *Lectures on Quantum Mechanics*. C.U.P., New York (2013)
- Wolfe, S.: On a continuous analogue of the stochastic difference equation $X_n = \rho X_{n-1} + B_n$. *Stoch. Processes Appl.* **12**, 301–312 (1982)
- Wong, E.: The construction of a class of stationary Markoff processes. In: *Stochastic Processes in Mathematics, Physics and Engineering*, Proc. Symp Appl. Math. XVI. American Mathematical Society, Providence, (1984)

Publisher's Note

Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.

Submit your manuscript to a SpringerOpen[®] journal and benefit from:

- Convenient online submission
- Rigorous peer review
- Open access: articles freely available online
- High visibility within the field
- Retaining the copyright to your article

Submit your next manuscript at ► [springeropen.com](https://www.springeropen.com)
